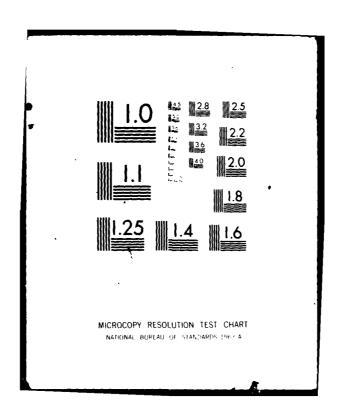
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ANALYSIS OF IOWA DATA I: INITIAL STUDY AND FINDINGS

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AND

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APRIL, 1980

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ANALYSIS OF IOWA DATA I: INITIAL STUDY AND FINDINGS

ABSTRACT

The first step of the analysis of Iowa Data with respect to the eventual application of the various new methods in latent trait theory is here initiated. The data are a set of approximately 500 item responses of each of 7,439 examinees to the Iowa Tests of Basic Skills, Form 6, on one of three difficulty levels, which correspond to the ages of 11, 12 and 13. The eleven subtests of this battery cover such areas as vocabulary, reading comprehension, language, mathematics and work-study skills, and have multiple choice formats, with 4 alternatives, except for one subtest in which there are five. After the initial tabulation, which includes various frequency distributions and the evaluation of its results, the elimination of certain examinees, who obviously skipped excessive numbers of items, is decided. An important question is whether the examinees' behavior fits the Knowledge and Random Guessing Principle, and then the three-parameter normal ogive or logistic model. The Chi-Square Goodness of Fit Test provides surprisingly little support. Index k*, which was introduced on an earlier stage of this research, was also used for this purpose.

The research was conducted at the principal investigator's laboratory, 409 Austin Peay Hall, Department of Psychology, University of Tennessee, Knoxville, Tennessee. Those who worked in the laboratory and helped the authors in various ways for this research include Paul S. Changas, Pamela Welch, Dete Furlan, C. I. Bonnie Chen, Philip Livingston, Linda Lance, and Melanie Perkins.

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The authors are obliged to Professor William Coffman of the University of Iowa and Director of the Iowa Testing Programs for providing us with valuable data for our research.

A special word of appreciation is due to Dr. Charles Davis of the Office of Naval Research, who arranged the initial contact between the principal investigator and Professor William Coffman.

The work was divided between the two authors in such a way that the senior author designed the research, introduced new concepts such as Informative Distractor Model and Equivalent Distractor Model, and so forth, and the junior author engaged in the actual work of data analysis. The bibliography on the three-parameter logistic model in Appendix I was made by the junior author.

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I Introduction

There have been many research projects in the history of psychology, in which researchers attempt to measure mental abilities by collecting data through the administration of a battery of tests. Even before the American Psychological Association adopted the concepts of reliability and validity coefficients in 1954, classical test theory, with all its variations, has practically the sole theoretical foundation for those research projects, and for a long time it has been the theory of mental measurement.

There are many deficiencies and limitations in classical mental test theory, however. The weakest points in classical mental test theory may be: 1) it heavily depends upon a specific population, or group, of examinees, and defines the characteristics of a test, or a test item, in relation with a specific group of examinees by naming them as if they were sole properties of a test, or a test item, while in reality they are the products of the interaction between the test, or the test item, and the group of examinees; and 2) it uses a single measure, the correlation coefficient, for its important concepts such as the reliabilty and validity of a test and treats it as if it were a magic number, while the fact is that many complexities are ignored and distorted because of its use. Thus many researchers have misled and lost themselves in the cobweb of classical mental test theory without being aware of the fact that they have produced nothing but

artifacts (cf. Samejima, 1977a).

Modern mental test theory, or latent trait theory, on the other hand, is based upon sounder axioms and rationale, and has gradually been invading the realm of classical mental test theory in the past decade, in the area of applied mental measurement. A relatively few, more sophisticated researchers have started conducting their research projects on mental measurement which is based upon latent trait theory. Among them, Rasch model (Rasch, 1960) may be the most popular model, because of its mathematical simplicity and easiness. When they deal with data in which multiple-choice test items are used, however, most researchers have turned to the three-parameter normal ogive, or logistic, model (Birnbaum, 1968), which is based upon the knowledge or random guessing principle in the context of latent trait theory. It is assumed that the examinee either knows the answer to a given multiple-choice test item, or guesses randomly, in selecting one of the given alternative answers. Thus the conditional probability, given ability, with which the examinee answers the item correctly is greater than the one with which the examinee knows the answer. When we consider this conditional probabilty as a function of ability, it is called the item characteristic function (Lord and Novick, 1968). In the three-parameter normal ogive, or logistic, model, this item characteristic function for a multiple-choice item is strictly increasing in ability, but its slope is less than the one for the item characteristic function in the normal ogive, or logistic, model for the free-response

test item. A bibliography of applied research based upon the three-parameter models is presented in Appendix I, together with one for theoretical works concerning the three-parameter models, which consists of papers selected, mainly, from recent issues of leading journals.

In spite of the popularity of the three-parameter logistic model, except for Lord (Lord, 1970), the researchers rarely have tried to validate the model in relation to their own data; they simply adopted the model, and assumed its validity. Such attitudes of researchers cannot be considered as being scientific, and it is imperative that more people should turn their attention to model validation, or to the search for a suitable model, or models, before they adopt one for their research. It was one of her considerations when the principal investigator started the series of research on theory and method of estimating the operating characteristics without assuming any mathematical form several years ago (Samejima, 1977a, 1977b, 1978a, 1978b, 1978c, 1978d, 1978e, 1978f). She has also proposed a new family of models for the multiple-choice test items (Samejima, 1979, 1980), in which she treats the multiplechoice item as something more than a blurred image of the free-response test item, which the three-parameter normal ogive, or logistic, model presumes. Thus in these new models each alternative wrong answer is considered to be a valuable information source, in addition to the correct answer.

In the present study of mental measurement based upon the

analysis of the Iowa Test data, unlike most research projects, we shall avoid setting any prior mathematical models, and start with the careful examination of our data, in order to determine which direction of research we should take. In other words, we are intentionally taking a slow process, and by no means are we anxious to extract hasty conclusions. This may seem to be tedious, but for truly scientific purposes of research this is the fastest way, if we wish to conduct research without sacrificing our conscience and with the promise of fruitfulness.

We hypothesize two main directions as our choices. One is the direction which leads to Informative Distractor Model, and the other is the direction which leads to Equivalent Distractor Model.

Note, however, that they are not specific mathematical models, but a very general categorization of ideas concerning the behavior of the multiple-choice test item. If, for instance, evidence clearly indicates the former direction, then we shall follow that direction to investigate the behavior of each test item more specifically. If the latter direction proved to be true, then we shall head for that direction and investigate the specifics about each test item. If it turns out that no single, general direction is indicated, then we shall depend upon the strategy of adopting theory and method which enable us to deal with test items following either general direction.

The present research is only the beginning of the analysis of the Iowa Test data, and more research is needed to add to it in the near future.

II Tests

The battery of tests used here is the <u>Iowa Tests of Basic</u>

<u>Skills, Form 6, Levels 9-14</u>. These tests have been designed,
constructed, and revised at the College of Education of the

University of Iowa since 1935, with the general school population
in mind, and for students of ages nine through fourteen, or grades
three through nine. All technical information in this paper has
been taken from either Form 6 itself (Hieronymous and Lindquist, 1971),
or its <u>Teacher's Manual</u> (Iowa Basic Skills Testing Program, 1971).

There are eleven tests in the battery, each of which focuses on a different basic skill. For convenience, hereafter, we shall call these separate tests subtests, in order to avoid the confusion which might occur when we refer to both the total test battery and each test in the battery. Following the Teacher's Manual, the descriptions and abbreviations of these eleven subtests, together with their administration schedule and working times, are tabulated and presented in Table 2-1. All the test items are power test items with multiple-choice format, with five alternative answers for the items in Subtest Ll, and with four alternatives for those in the other ten subtests. Within each subtest, test items are arranged in the ascending order of difficulty. These eleven subtests are designed to cover all major areas of academic interest for the grades three through nine. The separate directions for these eleven subtests are presented in Appendix II.

TABLE 2-1

Administration Sessions, Time Limits and Subtests of Iowa Tests of Basic Skills.

Administration Session	Working Time (Minutes)		Subtest
First Session 85 Minutes	17 55	V: V R: R	Vocabulary Reading Comprehension
Second Session 80 Minutes	12 15 20 20	L-1: L-2: L-3: L-4:	Spelling Capitalization Punctuation Usage
Third Session 85 Minutes	30 20 30	W-1: W-2: W-3:	Map Reading Reading Graphs and Tables Knowledge and Use of Reference Materials
Pourth Session 65 Minutes	30	M-1: M-2:	Mathematics Concepts Mathematics Problem Solving

The numbers of test items contained by the eleven separate subtests are 114, 178, 114, 102, 102, 86, 89, 74, 141, 136 and 96, respectively, following the order of subtests given in Table 2-1. They are all presented in a single test booklet of ninetysix pages. For each of the five levels, 9 through 14, only a subset of each subtest is administered. For example, for Subtest V, items 1 through 31 are given as Level 9, items 11 through 48 are administered as Level 10, items 24 through 66 are given as Level 11, and so forth. Different answer forms for the computerized scoring are provided for the six different levels. The standardized administration schedule and the working time for each subtest are presented in Table 2-1. For the entire test battery, the time required for the administration of each level of test is four hours and thirty-nine minutes. It is recommended that the test be administered on four consecutive days. It may occasionally have been presented on four consecutive half-days, but never within one day.

There are three options available to the classroom teachers in the administration of the test. Among them, graded testing is the most common one, in which a single level of test is administered to all the students in a given classroom. In this situation, Level 9 is given to the third graders, Level 10 to the fourth graders, Level 11 to the fifth graders, Level 12 to the sixth graders, Level 13 to the seventh graders, and Level 14 to both the eighth and ninth graders. The second option is

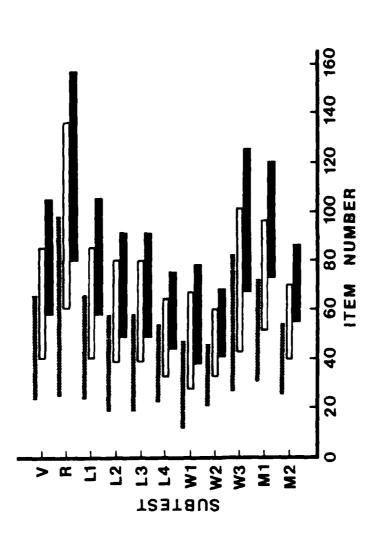
out-of-level testing. In this situation, the teacher selects an optimal level for his or her class, and administers that level of test, regardless of the examinees' nominal grade. The third option is individualized testing, in which each student in the class is given the level of test which is most suitable for his or her level of development. We notice that, even in this third situation, it is possible to administer the test as a group test with a single supervisor, since the administration schedule is standardized for all levels, as we can see in Table 2-1.

In our data, only the tests of Levels 11, 12 and 13 were used. In most cases, the first form of administration, i.e., graded testing, was adopted. There are a few exceptions, however, in which the third form was used. The numbers of test items contained in these three levels of test are 461, 487 and 500, respectively. Table 2-2 presents the number of test items in each of the eleven subtests, for each of the three levels. A graphical representation is made in Figure 2-1, to show how these three subsets of test items in each subtest overlap among the three levels. The numbers of test items in each subtest, which are included in all the three levels, in two adjacent levels, and in single levels, are shown in Table 2-3.

TABLE 2-2

Number of Items in Each of the Eleven Subtests and in Total for Each of the Three Levels, 11, 12 and 13.

Subtest Level	Λ	æ	TI	77	L3	77	WI	W2	E3	포	M2	Total
11	43	74 43	"	07	40	32	36	26	26 56	42	29	197
12	97	92	46	42	42	32	40	28	59	45	31	487
13	87	78	87	43	43	32	41	28	59	87	32	200



Test Items of Each of the Eleven Subtests Administered to Fach of Levels 11, 12 and 13, Which Are Represented by Shaded, Hollow, and Solid Bars, Respectively.

FIGURE 2-1

TABLE 2-3

Frequency Distribution of Items Which Were Administered to More Than One Level, As Well As Those Which Were Administered to One Level, and Itevel Only.

	•					
Level	11,12,13	11,12	12,13	11 Only	13 Only	Total
۸	6	27	28	16	20	100
æ	19	38	57	36	21	171
11	5	27	28	16	20	100
77	10	20	32	20	11	93
1.3	10	20	32	20	11	93
2	11	22	21	10	11	75
TA.	10	20	30	16	11	87
M2	9	14	20	12	∞	09
W3	16	07	35	16	24	131
Œ	0	21	54	21	24	06
M2	0	15	16	14	16	61
Total	100	797	323	197	177	1,061

III Data and Their Basic Tabulations

Data were collected in three different school systems in the State of Iowa, in the years 1971 through 1977. In their original form, the total number of examinees, including both boys and girls, is 7,581. Out of these people, 28 students took Level 9 Test and 114 took Level 10 Test. Since these are relatively small numbers, we decided to exclude them from our original group of examinees. The other 7,439 examinees are classified into three subgroups, i.e., 2,460 students who took Level 11 Test, 2,452 who took Level 12 Test, and 2,527 who took Level 13 Test. Hereafter, we shall call observations concerning these 7,439 examinees the original data.

Table 3-1 presents the frequency distribution of the test items for each of the eleven subtests with respect to the percentage of examinees who answered correctly, for each of Levels 11, 12 and 13. We can see that the configurations of these eleven frequency distributions are alike across the three levels, and, except for Subtest L1 for Level 13 and Subtest M2 for Levels 12 and 13, the medians of these frequency distributions are somewhat higher than 50 percent.

In these frequency distributions, all the no responses were treated as incorrect answers. Table 3-2 presents the frequency distributions of test items for each of the eleven subtests with respect to the percentage of examinees who answered in one way or another, for each of the three levels. We notice that, while

TABLE 3-1

Frequency Distribution of Items for Each of the Eleven Subtests with Respect to the Percentage of Examinees Answering Correctly. Each Interval of Percentage Is Greater than or Equal to the Lower End and Less than the Upper End. (Original Data)

Level 11

TABLE 3-1 (Continued): Level 12.

Percentage
7
9
2
7
7
œ
4
S
7
7
1
94

TABLE 3-1 (Continued): Level 13.

Total	0 0 1 1 1 1 1 2 0 0 0 0 0 0 0 0 0 0 0 0	200
M2	450 ESSET H 2H	32
M1	1 564556737	84
W3	10 10 3 3 3 3	29
W2	11 28282 242811	78
Wl	. 112222366322	41
Subtest 3 L4	7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7	32
Sul L3	7 5 3 6 7 7 3 4 3 1 5 5 1	43
1.2	132222221	43
L1	11 2 4 4 4 6 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	48
æ	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	78
>	10 10 10 10 7	84
Percentage	0.0 - 5.0 5.0 - 10.0 10.0 - 15.0 15.0 - 20.0 20.0 - 25.0 25.0 - 30.0 36.0 - 45.0 45.0 - 45.0 45.0 - 45.0 45.0 - 60.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 80.0 - 95.0 90.0 - 95.0	Total
	10 8 7 8 10 11 11 11 11 11 11 11 11 11 11 11 11	

TABLE 3-2

ŗ

Frequency Distribution of Items for Each of the Eleven Subtests with Respect to the Percentage of Examinees Answering, Though Not Necessarily Correctly. Each Interval of Percentage Is Greater than the Lower End and Less than or Equal to the Upper End. (Original Data)

Level 11

99.5% 100.0%	15	33	12	16	15	œ	6	11	7	13	9	142
99.0- 99.5%	5	7	6	5	6	15	7	<u>س</u>	15	&	9	89
99- 100%	20	40	21	21	24	23	16	14	19	21	12	231
98- 99%	8	7	5	4	7	9	2	4	10	7	2	89
97-	2	7	٣	5	5	3	က	2	2	٦	Н	34
-96 -21%	2	-	2	2	7		-	-	3	5	2	23
95-	2	ო	-	m			-	7	7	7	7	16
94-	2	3	2	Н			Н	1	2	-	-	14
93-	2	4	Н	7			Н	2	-	2	1	15
92- 93%		H	Н	2				Н	7		7	8
91- 92%	3	2	7	7			Н			2		10
90- 91%	H	2	H						2	-1	7	8
-0 80%	1	4	2				7		14		က	34
Percentage Subtest	Λ	œ	L1	L2	L3		WI	W2	M3	M	M2	Total

TABLE 3-2 (Continued): Level 12.

99.0- 99.5- 99.5% 100.0%	5 26	8 37	3 23	3 25	3 31	7 24	5 16	2 16	7 31	11 17	11 8	65 254
200%	31	45	26	28	34	31	21	18	38	28	19	6]
99% 100%	٠	7	7	2	(,)	(*)	7		(,,		1	319
-86 99%	2	2	7	7	œ	1	2	2	2	4	3	43
97-	2	2	2	7			-	7	4	2	2	25
96-	2	2	2	3			3	7	7	က	1	23
95- 96%	2	7	2	က			7		-	7	1	13
94-	3	2	-	2			-	~	က	7	1	16
93- 94%	1	4	H				-		2	2	1	12
92- 93%		3	1					H	2		-	8
91- 92%		7	1				-1	Н	2		-	8
90-		4						1				5
-0 80%			9				80				-4	15
Percentage Subtest	Λ	œ	L1	L2	L3	7.7	WI	W2	W3	M	M2	Total

TABLE 3-2 (Continued): Level 13.

99.0- 99.5- 99.5% 100.0%	27	36	25	27	34	28	18	17	31	29	14	286 0
99- 100%	27	36	25	27	34	28	1.8	17	31	29	14	286
%66 86-	11	21	5	9	6	4	9	4	11	6	6	95
97-	9	4	4	2			7	2	4	2	3	34
-96 -96	7	9	٣	2			2	2	3	3	1	29
95-		7						1	က	2		11
94- 95%		7	2				2	2	3	က	1	20
93-			2						3		-	9
92-			2				1		-		2	9
91-							2				-	3
90-			3				1					7
%06 -0			~				5					9
Percentage Subtest	Λ	~	L1	1.2	1.3	174	W1	W2	W3	M1	M2	Total

the majority of items, i.e., 231 out of 461 for Level 11, 319 out of 487 for Level 12, and 286 out of 500 for Level 13, have attracted some answers from 99 percent or more of the examinees, there are some other items, i.e., 34 for Level 11, 15 for Level 12, and 6 for Level 13, to which only less than 90 percent of the examinees have responded. The cumulative frequency of test items to which one of the six given percentages, or more, of the examinees responded is presented in Table 3-3, for each of the eleven subtests, and for each of the three levels.

The breakdown of the last column of Table 3-2 into two subcategories, i.e., 99.0 to 99.5 percent and 99.5 to 100.00 percent, which are added to its right as two additional columns of Table 3-2, has revealed that, for Level 13, there are no items to which 99.5 percent, or more, of the examinees responded. This result looks rather strange, and has urged us to shift our attention back to our raw data. As the result, it was found out that there are a small number of examinees who did not respond to a substantially large number of test items. Table 3-4 presents the frequency distributions of examinees with respect to the number of unanswered test items for Levels 11, 12 and 13. We can see in this table that, while as many as 7,010 examinees out of 7,439 left only 49 or less test items unanswered, there also are 162 examinees who did not respond to as many as 100, or more, test items. Our raw data show there are some examinees included who skipped an entire subtest, or more than one entire subtest.

TABLE 3-3

Frequency Distribution of Items Which Were Answered, Though Not Necessarily Correctly, by Specified Percentages of Examinees or More. (Original Data)

Level 11

Subtest	Number of Items	Min 90%	imum 1	Percei		Resp 96%	
		-					
V	43	42	38	36	34	32	30
R	74	70	66	61	58	55	54
Ll	43	38	36	34	32	31	29
12	40	40	39	36	35	32	30
L3	40	40	40	40	40	40	36
L4	32	32	32	32	32	32	32
W1	36	29	28	27	26	25	24
W2	26	26	26	23	22	21	20
w3	56	42	40	37	35	34	31
MI	42	42	39	37	36	34	29
M2	29	26	25	23	22	20	18
Total	461	427	409	386	372	356	333

TABLE 3-3 (Continued): Level 12.

Subtest	Number of Items	Min 90%	imum 1		•	Respo	97%
v	46	46	46	45	42	40	38
R	76	76	70	63	61	60	5 5
Ll	46	40	39	37	36	34	32
L.2	42	42	42	42	40	37	34
L3	42	42	42	42	42	42	42
L4	32	32	32	32	32	32	32
Wl	40	32	31	30	29	27	24
W2	28	28	26	25	24	24	22
w3	59	59	57	53	50	49	47
Ml	45	45	45	43	41	40	37
M2	31	30	29	27	26	25	24
Total	487	472	459	439	423	410	387

TABLE 3-3 (Continued): Level 13.

Subtest	Number of Items	Min 90%	imum 1		ntage 95%	Resp	
v	48	48	48	48	48	48	44
R	78	78	78	78	71	67	61
L1	48	47	44	40	38	37	34
L2	43	43	43	43	43	43	38
L3	43	43	43	43	43	43	43
L4	32	32	32	32	32	32	32
W1	41	36	33	32	30	30	28
W2	28	28	28	28	26	25	23
w3	59	59	59	55	52	49	46
MI	48	48	48	48	45	43	40
M2	32	32	31	28	27	27	26
Total	500	494	487	475	455	444	415

TABLE 3-4

Frequency Distribution of Examinees with Respect to the Number of Items Which Were Left Unanswered for Each of the Three Levels, 11, 12, and 13. (Original Data)

Number Unanswered	11	Leve 12	1 13	Total
0-49	2271	2326	2413	7010
50-99	120	85	62	267
100-199	55	34	29	118
200-299	13	7	18	38
300-500	1	0	5	6
Total	2460	2452	2527	7439

Table 3-5 presents the frequency distribution of such examinees with respect to the number of subtests which were omitted entirely, for each of the three levels. We can see in this table that one examinee who did not participate in the testing at all is included in our group of examinees. The relationship between the omission of one or more entire subtests and the number of unanswered items is shown in Table 3-6 in the form of the frequency distribution of these 67 examinees with respect to the level and the number of unanswered test items. Comparison of this table with Table 3-4 indicates that the six examinees who left 300 or more test items unanswered are among those who omitted, at least, one entire subtest, and as many as 27 out of 38 examinees who omitted 200 to 299 test items also belong to this subgroup of examinees. A frequency distribution similar to Table 3-6 was made for the 126 examinees who left, at least, one half of a subtest, but not the entire subtest, unanswered, and is presented as Table 3-7. This additional information indicates that, if we exclude all the examinees who left, at least, one half of a subtest unanswered from our total group of examinees, then the number of examinees who left 200 or more test items unanswered will become zero, and only 28 examinees, who omitted more than 100, but less than 200, test items, will be included. For this reason, we have decided to exclude these 193 examinees from our original group of examinees for further analysis. Hereafter, we shall call observations concerning the remaining 7,246 examinees the revised data, to

TABLE 3-5

Frequency Distribution of Examinees Having Omitted, at Least, One Entire Subtest with Respect to the Total Number of Unanswered Subtests for Each of the Three Levels, 11, 12, and 13. (Original Data)

[eve]	·· ··· ·····	z	Number of Subtests Omitted Entirely	of	Subtes	its Om	itted	Enti	rely			F 0 4 0
	-	2	3 4 5 6 7	7	2	9	7	8	9 10 11	10	11	IOLAI
11	7	ო	1	က	1	2	3				Н	15
12	7	2		-	2							10
13	3	3 10 7	7	2	8 10	10	i	-	-			42
Total 6 18 8 6 11 12 3 1 1 0 1	9	18	8	9	11	12	3	1		0	-	67

TABLE 3-6

Frequency Distribution of Examinees Who Omitted, at Least, One Entire Subtest with Respect to the Total Number of Items Omitted, for Each of the Three Levels, 11, 12, and 13. (Original Data)

Number Unanswered	11	Leve:	13	Total
0-49	0	0	0	0
5C-99	2	5	5	12
100-199	6	2	14	22
200-299	6	3	18	27
300-500	1	0	5	6
Total	15	10	42	67

TABLE 3-7

Frequency Distribution of the Examinees Having Left, at Least, One Half of a Subtest Unanswered, But Not Omitted Any Entire Subtest, with Respect to the Number of Unanswered Items, for Each of the Three Levels, 11, 12, and 13. (Original Data)

Number Unanswered	11	Leve:	13	Total
0-49	9	1	3	13
50-99	25	5	4	34
100-199	40	19	9	68
200-299	7	4	0	11
300-500	0	0	0	О
Total	81	29	16	126

distinguish themselves from the original data.

Table 3-8 presents the item identifications of the fifty-five test items, i.e., 34 for Level 11, 15 for Level 12, and 6 for Level 13, to which less than 90 percent of examinees responded in the original data, the percentages of examinees who answered in one way or another in the original data, and those in the revised data. We can see in this table that for most of these fifty-five test items the two percentages show a visible improvement caused by the exclusion of the 193 examinees. The frequency distributions of test items for the eleven subtests with respect to the percentage of examinees who answered in one way or another in the revised data are presented in Table 3-9, for each of the three levels. Comparison of this table with Table 3-2 reveals a substantial improvement in the percentage provided by the exclusion of the 193 examinees for all the three levels. Among others, we notice that the frequency of test items which were answered by 99 percent, or more, of examinees increased from 231 to 320 for Level 11, from 319 to 350 for Level 12, and from 286 to 377 for Level 13.

Table 3-10 presents the frequency distribution of test items for each of the eleven subtests with respect to the percentage of examinees who answered correctly, for each of Levels 11, 12 and 13, after the exclusion of the 193 examinees. There are a substantial number of categories whose frequencies changed from those in Table 3-1, although the two frequency distributions for each

TABLE 3-8

Fifty-five Test Items to Which Less Than Ninety Percent of Examinees Responded in One of the Three Levels in the Original Data, the Percentages of Responses in the Original Data, and Those in the Revised Data.

	Leve	1 11	Leve	1 12	Leve	1 13
It-m	Original	Revised	Original	Revised	Original	Revised
V-66	89.1	91.4				
R-95	89.6	91.9	,			
R-96	89.2	91.5				
R-97	1 88. 7	91.0				
R-98	88.3	90.7				
L1-62	89.4	91.8	1		1	
L1-63	88.3 87.5	90.7 90.0	!			
L1-64 L1-65	86.3	88.7				
L1-66	84.8	87.3				
L1-80			89.7	90.7	:	
L1-81	1		88.9	89.9		
L1-82			87.9	88.9	٠ .	
L1-83			87.0	88.0	:	
L1-84	1 ,		86.2	87.2		
L1-85	1		85.4	86.4	į	
L1-105	! 1	01 (<u>;</u>		89.7	90.6
W1-41	88.9	91.4 88.2			. !	
W1-42 W1-43	83.3	86.0			i.	
W1-43 W1-44	81.7	84.3	i t		. 1	
W1-45	79.2	81.6	1			
W1-46	76.7	79.0			į l	
W1-47	74.7	76.9	; 1			
W1-60	1		89.2	90.4	· .	
W1-61	i		87.2	88.3	•	
W1-62	1		85.2	86.3		
W1-63	!		82.7	83.8	· .	
W1-64			80.8	81.9		
W1-65	l i		78.4	79.4		
W1-66 W1-67	1		75.4	76.3	• :	
W1-74			74.2	75.2		
W1-75	· .		i i		88.9	90.0
W1-76			,		87.3 86.2	88.4 87.2
W1-77	<u> </u>		!		85.0	86.1
W1-78	1		+		83.9	84.9
W3-69	90.0	92.3	1		,	04.5
W3-70	89.3	91.5	i i		:	
W3-71	88.8	91.0	! 1		1	
W3-72	87.9	90.2				
W3-73	87.1	89.4	!			
W3-74 W3-75	86.8	89.0				
W3-75 W3-76	86.0	98.3	, i		t	
W3-76 W3-77	84.0	87.2 86.3	i			
W3-78	83.5	85.8	k - 1			
W3-79	82.8	85.0	. 1		1 :	
W3-80	82.3	84.5	'		. '	,
W3-81	81.6	83.8	:			
W3-82	, 91.1	33.3	,			i I
M2-52	87.9	90.0				
M2-53	. 85.9	87.9				
M2-54 M2-69	93.7	85.7	20. 1			
.14 -07	•		. 39.4	90.3		

TABLE 3-9

Frequency Distribution of Items for Each of the Eleven Subtests with Respect to the Percentage of Examinees Answering, Though Not Necessarily Correctly, Each Interval of Percentage Is Greater than the Lower End and Less than or Equal to the Upper End.

Level 11

Percentage Subtest	0- 90%	90- 91%	91-	92- 93%	93-	94-	95-	96- 97%	97- 98%	-86 -86	18- 99- 99% 100%	Total
				1	2	1	2	2	2	7	28	43
	-	2	2	1	2	2	7	4	2	2	50	7.4
	т	-	=	1		Н	2	2	-	4	27	43
					-	1	2	-	3	3	29	40
										5	35	40
										~	31	32
	9		-		-		-	-	-	2	23	36
	-					1	-	2	-	-	20	26
	10	-	2	2	-	. 1	2	-	2	3	31	95
				1	2		2	٦	2	3	28	42
	~	ļ		-		-	1	-	2	2	18	29
	22	7	7	7	6	∞	17	15	19	33	320	461

TABLE 3-9 (Continued): Level 12.

												
Tota1	97	9/	97	42	42	32	07	28	59	45	31	187
8- 99- 99% 100%	34	87	29	31	38	32	23	20	43	32	20	350
98- 99%	7	7	3	က	4		2	-	7	2	0	36
97- 98%	2	2	2	m			2	Э	2	3	2	24
96- 97%	2	-	2	3			2		П	7	П	13
95-	3	2	-4	2			-	-	3	2	1	16
94-	1	4	7				H		3	2	1	13
93-		3	1					-	2			7
92- 93%		7	-				-	1	Н		2	10
91-		2						-			ļ	3
90-			-				-				1	3
206 -0			2				7					12
Percentage Subtest	>	œ	[]	1.2	L3	174	Z	w2	M3	æ	M2	Total

TABLE 3-9 (Continued): Level 13.

Total	87	78	87	43	43	32	41	28	65	87	32	200
8- 99- 99% 100%	38	99	30	33	42	32	24	21	42	37	22	377
0,	∞	9	4	2	1		4	2	4	3	4	41
97-	2	5	3	2			2	2	8	3		25
96-		7	1					-	3	2	7	12
95-		7	3				2	2	3	3	7	21
94-			-						3			5
93-			1				2		7		2	9
92-			1				-				1	2
91-			3									4
90-			7									-
%06 -0		-					2					5
Percentage Subtest	Λ	æ	1.1	1.2	L3	174	W1	W2	W3	M1	M2	Total

TABLE 3-10

Frequency Distribution of Items for Each of the Eleven Subtests with Respect to the Percentage of Examinees Answering Correctly. Each Interval of Percentage Is Greater than or Equal to the Lower End and Less than the Upper End.

Level 11

$\overline{}$	Percentage	>	æ	L1	L2	Sub L3	Subtest 3 L4	Wl	W2	W3	M1	M2	Total
-	ļ ,												
	5.0 - 10.0												> c
	1												o C
	ı		1	2	-								7
	ı		-	1	-		1	Н				٦,	7
	1	7	2				1	1				7	6
	1	_	7	2			Т	7		Н	m	7	16
	ı	8	2	2	H	7	Ţ	2		-	7	e	21
	1	4	2	9	É	7	5	٦	7	e	٣	2	38
	45.0 - 50.0	4	6	7	9	5	9	-	3	7	3		87
_	1	4	8	, (*)	٣		4	4	7	10	œ	3	54
	55.0 - 60.0	10	2	2	7	5	7	9	3	15	3	2	09
	1	2	4	-	6	5	5	4	2	9	2	3	97
	65.0 - 70.0	7	ò	က	2	6	٣	9	n	6	2	7	57
	1	7	6	S	e	7	7	2	1	9	٣	ĸ	39
	t	2	2	9	4	9		~	1	-	7	7	32
	ł	-1	7		-	-		7	2		Н		15
	1		4		1	-		7	7		-		6
_	ı		7					2	~		H		9
	95.0 -100.0											_	0
		:	;	1:	:] :		}			1	!	
	Total	43	74	43	40	40	32	36	26	99	42	59	461

TABLE 3-10 (Continued): Level 12.

Total	000	. H &	15	20	29	52	20	61	24	43	47	38	21	6	2	0	487
M2		p-4	2	7	. 6	2		2	_		5	3					31
Ml		2	2	п с	7	7	7	4	4	7	m	7	7	7			4.5
W3			1	ღ 4	7	2	2	9	H	1.0	80	2	∞	2			59
W2					7	1	П	7	9	2	7	2	1	1			28
W1			1	m <	7	7	4	7	7	7	7	7		1			07
Subtest 3 L4			2	2 5	1	9	2	2	2	3	-						32
Sub L3				~	· ~	9	∞	9	4	2	2	7		П			42
1.2		~		2	m	Н	2	2	9	5	5	9	4		-4		42
=		H	S	יט יע	5	7	7	9	2	7	-	2	7				46
24		H	П.	e 4	2	10	12	6	12	2	9	7	2	2	-		9/
Λ		H		T 4		3	4	7	∞	7	2	7	2				46
Percentage	0.0 - 5.0 5.0 - 10.0		ſ	1 1	40.0 - 45.0	1	1	1	1	i	t	- C.	1	t	1	95.0 -100.0	Total
	1 2 3 3	, 4 N	9 1	~ 8	6	10	11	12	13	14	15	91	17	18	19	20	

32 M22 2 1 4 4 2 87 M 59 W3 28 32233 WZ 1122333762311 4.1 Z Subtest L3 L4 32 3245413321 33677423122 43 1447336425151 43 L2 48 Γ 1 2 3 4 4 4 6 6 6 7 7 7 7 7 7 15 5 7 \simeq 78 > 222446662444 48 0.0 - 5.0 5.0 - 10.0 10.0 - 15.0 15.0 - 20.0 20.0 - 25.0 25.0 - 30.0 30.0 - 35.0 40.0 - 45.0 45.0 - 50.0 50.0 - 55.0 50.0 - 55.0 50.0 - 55.0 50.0 - 60.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 60.0 - 65.0 Percentage

TABLE 3-10 (Continued): Level 13.

Total

-35**-** III-5

subtest and each of the three levels are similar in configuration.

It should be noted that, even in the revised data, these percentages correct are not independent from the positions of the test items in each subtest. The relative frequencies of examinees who left the items unanswered are given in Appendix III, in which the test items are arranged in the order of presentation within each subtest, for each of the three levels. There is a distinct tendency that larger numbers of examinees did not respond to items which were presented later in each subtest. It is obvious, therefore, that, for these later items, the percentage for the correct answer is less than it should be in the ideally set free-response situation.

IV Informative Distractor Model or Equivalent Distractor Model?

By <u>Informative Distractor Model</u>, we mean the family of models in which we assume the existence of specific information obtainable from separate alternative answers, including the correct answer, of each multiple-choice test item. The family of models, including Models A, B and C, proposed by Samejima (Samejima, 1979) belongs to this general model.

In contrast to this, by Equivalent Distractor Model we mean the family of models in which no specific information is expected from separate incorrect answers, which are given as alternatives in the multiple-choice test item. Thus all the alternatives, except for the correct answer, of a given multiplechoice item are equivalent, since the information given by a specific alternative, or distractor, is not different from the one given by any one of the remaining wrong answers. The threeparameter normal ogive, or logistic, model belongs to this family of models. In this model, all the information provided by a given wrong answer is pure noise resulting from random guessing, and, therefore, the alternative is equivalent with any one of the remaining wrong answers. Note, however, that this type of model is not the only one included by the Equivalent Distractor Model. Suppose that the operating characteristic of each wrong answer of a given multiple-choice item includes some information about the examinee's ability, but all the operating characteristics, or plausibility curves, of the distractors are

identical. In such a case, we can say that the item should belong to the Informative Distractor Model in the sense that these distractors provide us with some information concerning the examinee's ability. On the other hand, we can also say that the item should belong to the Equivalent Distractor Model, since each distractor does not have any specific information which distinguishes itself from the other distractors. For convenience, in the present paper, we shall take the second standpoint, defining the Informative Distractor Model in the narrower sense.

As an indicator which enables us to determine the direction each test item should take, we shall use Index k*, which was proposed by Samejima (Samejima, 1980). This index was originally developed for the purpose of invalidating the knowledge or random guessing principle, upon which the three-parameter normal ogive, or logistic, model is based. It is also useful as the first step of searching the direction for each test item, with the two general models in mind.

Let \overline{A} be the event that the examinee <u>does not</u> know the answer to the multiple-choice item g. We shall consider the probability space which consists of the subpopulation of such examinees. The conditional probability, $p(i|\overline{A})$, with which the examinee selects the alternative $i(=1,2,\ldots,mg$, or m) of item g in this conditional probability space is given by

(4.1)
$$p(i|\overline{A}) \begin{cases} = p_{i} \begin{bmatrix} \sum p_{i} + p_{R}^{*} \end{bmatrix}^{-1} & i \neq R \\ i \neq R & i \neq R \end{cases}$$

$$= p_{R}^{*} \begin{bmatrix} \sum p_{i} + p_{R}^{*} \end{bmatrix}^{-1} , \quad i = R$$

where p_i is the probability with which the examinee chooses the alternative i, and p_R^{\star} is the probability with which the examinee guesses correctly for item g. Index k^{\star} is defined in terms of these conditional probabilities, such that

(4.2)
$$k^* = \exp\left[-\sum_{i=1}^{m} p(i|\overline{A}) \cdot \log p(i|\overline{A})\right] = \left[\prod_{i=1}^{m} p(i|\overline{A})^{p(i|\overline{A})}\right]^{-1}.$$

It is obvious that this conditional probability, $p(i|\overline{A})$, for a wrong answer, or $i\neq R$, is proportional to p_i , since, according to the knowledge or random guessing principle, every examinee in the original population who has selected a wrong answer does not know the correct answer to item g, and, consequently, belongs to the subpopulation \overline{A} . On the other hand, we can write

$$(4.3) p_R^* \leq p_R^*,$$

for examinees who have selected the correct answer R do not necessarily belong to the subpopulation \overline{A} .

Let θ be ability, or latent trait, and $P_g(\theta)$ be the item characteristic function of the multiple-choice item g. In the three-parameter normal ogive, or logistic, model, this item characteristic function is given by

(4.4)
$$P_{g}(\theta) = \Psi_{g}(\theta) + [1-\Psi_{g}(\theta)]c_{g} = c_{g} + [1-c_{g}]\Psi_{g}(\theta),$$

where $\Psi_{\mathbf{g}}(\theta)$ is the item characteristic function of item \mathbf{g} when it is given as a free-response test item, which is specified by

(4.5)
$$\Psi_{g}(\theta) = (2\pi)^{-1/2} \int_{-\infty}^{a_{g}(\theta-b_{g})} e^{-u^{2}/2} du$$

in the normal ogive model, and by

(4.6)
$$\Psi_{g}(\theta) = [1 + \exp\{-Da_{g}(\theta - b_{g})\}]^{-1}$$

in the logistic model, and c_g is the guessing parameter which equals 1/m. Let $f(\theta)$ denote the density function of ability θ , and p_R and p_g be the probabilities with which the examinee answers the item correctly in the multiple-choice situation and the free-response situation, respectively. Thus we can write

(4.7)
$$p_{g} = \int_{-\infty}^{\infty} \Psi_{g}(\theta) f(\theta) d\theta ,$$

and

(4.8)
$$p_{R} = \int_{-\infty}^{\infty} P_{g}(\theta) f(\theta) d\theta = p_{g} + c_{g}(1-p_{g}) .$$

It is noted that the second term of the rightest hand side of (4.8) is the probability with which the examinee guesses correctly, i.e.,

 p_R^{\star} . We also notice that this quantity equals the probability with which the examinee chooses a distractor i $(\not=R)$. Thus we can write

(4.9)
$$p_i = c_g(1-p_g) = p_R^*$$
 ($i\neq R$)

It is obvious from (4.9), (4.1) and (4.2) that Index k* assumes its maximal value, m, when the three-parameter normal ogive, or logistic, model holds. Note, however, that it is a necessary condition for the validity of the model, but not a sufficient condition. Index k* can be used for the invalidation of the three-parameter normal ogive, or logistic, model, therefore, but not for the validation of the model, unless it is combined with other evidence.

In practice, we obtain the estimate of Index k*, by replacing $p(i|\overline{A})$ in (4.2) by its estimate, $\hat{p}(i|\overline{A})$. To obtain this estimate, we can use the frequency ratio, P_i , as the estimate of the probability, p_i , for each wrong answer, or distractor, i ($\neq R$). We notice, however, that the estimate, P_R^* , for the probability p_R^* , is not directly observable from our data, and, in one way or another, must be defined indirectly. In congruence with the purpose of invalidating the knowledge or random guessing principle, such a strategy is taken that we find P_R^* which makes the resulting estimate of Index k* maximal. In so doing, we define entropy \hat{H}^* such that

(4.10)
$$\hat{H}^* = \log \hat{k}^* = -\sum_{i=1}^{m} \hat{p}(i|\overline{A}) \cdot \log \hat{p}(i|\overline{A}) .$$

Defining P_i^* such that

(4.11)
$$P_{i}^{*} \begin{cases} = p_{i} & i \neq R \\ = P_{R}^{*} & i = R \end{cases},$$

we can write for the estimate, $\hat{p}(i|A)$, such that

(4.12)
$$\hat{p}(1|\overline{A}) = P_{1}^{*} \begin{bmatrix} m \\ \Sigma P_{1}^{*} \end{bmatrix}^{-1}.$$

Then we can rewrite (4.10) in such a way that

(4.13)
$$\hat{H}^* = -\begin{bmatrix} \Sigma P^* \end{bmatrix}^{-1} \begin{bmatrix} \Sigma P^* \cdot \log P^* - (\Sigma P^*) \cdot \log \{\Sigma P^* \} \end{bmatrix} .$$

$$s=1 \quad i=1 \quad s=1 \quad$$

We have for the partial derivative of $\,\hat{H}^{\star}\,\,$ with respect to $\,\,P_{R}^{\star}\,\,$ such that

(4.14)
$$\frac{\partial \hat{H}^{*}}{\partial P_{R}^{*}} = \begin{bmatrix} \sum_{s=1}^{m} P_{s}^{*} \end{bmatrix}^{-2} \begin{bmatrix} \sum_{i=1}^{m} P_{i}^{*} \cdot \log P_{i}^{*} - (\sum_{s=1}^{m} P_{s}^{*}) \cdot \log P_{R}^{*} \end{bmatrix},$$

and, setting this derivative equal to zero, we obtain

$$P_{R}^{\star} = \prod_{\substack{i \neq R}} P_{i} \left[\sum_{s \neq R} P_{s} \right]^{-1},$$

or

(4.16)
$$\log P_{R}^{\star} = \left[\begin{array}{c} \mathbb{Z} P_{s} \end{array} \right]^{-1} \quad \mathbb{Z} P_{i} \cdot \log P_{i} \quad .$$

It is obvious that Index k* introduced in the preceding paragraphs can be used for our purpose of searching the direction that each multiple-choice item should take, i.e., Equivalent Distractor Model or Informative Distractor Model. If Index k* turns out to be far less than m, then we must reject the hypothesis of Equivalent Distractor Model for that test item. If it assumes a value close to m, then we shall say that Equivalent Distractor Model should still be under consideration. In both cases, Informative Distractor Model stays among the possibilities.

It is noted that the traditional chi-square test with (m-2) degrees of freedom for the goodness of fit for the frequencies of the (m-1) wrong answers with the uniform distribution as the theoretical distribution may serve our purpose just as well, without using Index k*. In our pilot study, we applied it for the original data of 7,439 examinees, and the result is summarized in Appendix IV. Table A-4-1 presents the frequency distribution of test items for each of the eleven subtests with respect to the probability resultant from this chi-square test, for each of Levels 11, 12 and 13. As we can see in these tables, only 23, 22 and 21 test items indicate the acceptance of the respective uniform distributions, or the acceptance of Equivalent Distractor Model, for Levels 11, 12 and 13, respectively, even if we take as low a level of significance as 0.0005. Table A-4-2 presents the item identification, the probability and the percentage of the correct answer of each of these 66 test items. We can see

in this table that, if we raise the level of significance to 0.01, then all the items whose probabilities are marked with ** will be excluded, to make the total number of items as small as 45, i.e., 17 for Level 11, 11 for Level 12, and 17 for Level 13. If, further, we raise the level of significance up to 0.05, then the total number of items will be reduced to 36, i.e., 15 for Level 11, 7 for Level 12, and 14 for Level 13, excluding all the items whose probabilities are marked with * . This number is only 2.5 percent of the total number of items, 1,448. There are only 19 items whose probabilities are greater than 0.2, i.e., R-78, W1-21, W1-25, W3-39, M1-31, M1-38, M1-45 and M1-68 for Level 11, V-79, R-78, W1-29, W2-44, W3-53 and M1-95 for Level 12, and V-93, R-140, R-155, W2-44, and M1-107 for Level 13. Since it is unlikely that this result obtained upon the original data is substantially different from the one obtainable from our revised data, we must take it as the suggestion for the rejection of Equivalent Distractor Model. Note, however, that the chi-square test applied for our data has severe criteria, since, in most cases, our sample sizes are very large. It is interesting to note that, except for W1-25 (57.9%) of Level 11, V-79 (46.6%) and M1-95 (29.4%) of Level 12, and V-93 (49.9%), R-140 (33.8%), R-155 (35.7%) and M1-107 (40.6%) of Level 13, all the other 12 test items, whose probabilites are greater than 0.2, have greater percentages of the correct answer than the medians of the separate subtests, as is clear if we compare Table A-4-2 with Table 3-1.

Since the examinees who unswered correctly were excluded from those used for the chi-square test, we can see that these test items whose probabilities are greater than 0.2 are based upon relatively small sample sizes. As a whole, however, sample sizes are large, and the chi-square test is very sensitive to small deviations from the assigned uniform distributions.

In contrast to the chi-square test, we can see from (4.7) that the estimated Index k* is <u>insensitive</u> to the sample size, because the sampling fluctuation participates in the resulting estimate only through the computation of the proportions, P₁. Thus, whether it is right or wrong, if we wish to ignore the sampling fluctuations of the proportions, then the values of estimated Index k* can be comparable across different sample sizes.

Table 4-1 presents the frequency distribution of the items of each of the ten subtests, excluding Subtest L1, which consists of five-alternative test items, with respect to the resultant values of the estimated Index k*, for each of Levels 11, 12 and 13. The corresponding result for Subtest L1 is presented, separately, as Table 4-2, for all the three levels. We can see in Table 4-1 that the configurations of these frequencies are similar across the three levels, with the range of the estimated Index k*, 2.25 through 4.00, for each level. This is also the case with Subtest L1, with the range of estimated Index k* 2.25 through 4.50 for most items, as is shown in

TABLE →-.

Frequency Distribution of Four-Alternative Items with Respect to Index k* for book if the len Subtests. The Range of Index k* is reafer Than or iqual to the lower bind and Less Than the opper rind of Each Interval.

Level 11

	Range of Index k*	v	R	1.2	L3	Sub L4	test W1	W2	W3	M1	H2	Total
1	1.00 ~ 1.25								-			0
1	1.25 - 1.50											1 0
2		ŀ										1 0
3	1.50 - 1.75											"
•	1.75 - 2.00											0
5	2.00 - 2.25											0
6	2.25 - 2.50	l	1	2				1				4
7	2.50 ~ 2.75	1	2	7	1		ı			1		13
8	2.75 - 3.00	3	2	6	8	3		2	1			25
9	3.00 - 3.25	6	6	10	12	8	1	2	5	3	1	54
10	3.25 - 3.50	3	13	8	7	12	4	1	9	4	ì	62
11	3.50 - 3.75	11	13	6	8	6	7	4	12	7	12	86
12	3.75 - 4.00	19	37	1	4	3	23	16	29	27	15	174
	Total	43	74	40	40	32	36	26	56	42	29	418

Level 12

	Range				_	Subt	est					Total
	of Index k*	V	R	1.2	L3	1.4	W1	W2	W3	H1	H2	10021
1	1.00 - 1.25											0
2	1.25 - 1.50	}										0
1 3	1.50 - 1.75	1										0
4	1.74 - 2.00	}										0
1 5 1	2.00 - 2.25	ł										0
6	2.25 - 2.50	Į.	1	4								5
7	2.50 - 2.75	1 2	1	8	1	1				1		14
8	2.75 - 3.00	2	4	6	8	3			5	2		30
9	3.00 - 3.25	4	10	7	8	8	2		8	6	1	54
10	3.25 - 3.50	6	9	8	10	11	4	3	11	5	3	70
11	3.50 - 3.75	10	18	5	11	5	6	8	16	8	10	97
12	3.75 - 4.00	22	33	4	4	4	28	17	19	23	17	171
	Total	46	76	42	42	32	40	28	59	45	31	441

Level 13

	Range					Sub	test					
	of Index k*	V	R	L2	L3	14	W1	W2	W3	M1	H2	Total
1	1.00 - 1.25											0
2	1.25 - 1.50	}										0
3	1.50 - 1.75	Į.										0
4	1.75 - 2.00	i										0
5	2.00 - 2.25	, 1										0
6 1	2.25 - 2.50	2		3								5
, ,	2.50 - 2.75	1 3		7	2	1			1			14
8 '	2.75 - 3.00	2	5	7	4	2			2	2		24
9	1.00 - 1.25	1	5	10	11	7			7	4	1	46
10	3.25 - 3.50	11	7	8	10	10	5		10	9	5	75
11	3.50 - 3.75	10	24	5	11	6	7	10	21	12	7	113
12	3.75 - 4,00	19	37	3	5	6	29	18	18	21	19	175
	Total	48	78	43	43	32	41	28	59	48	32	452

TABLE 4-2

Frequency Distribution of Five-Alternative Items of Subtest L1, with Respect to Index k*, for Levels 11, 12, and 13, Respectively.

	Range of Index k*	11	Level	13	Total
1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16	1.00 - 1.25 1.25 - 1.50 1.50 - 1.75 1.75 - 2.00 2.00 - 2.25 2.25 - 2.50 2.50 - 2.75 2.75 - 3.00 3.00 - 3.25 3.25 - 3.50 3.50 - 3.75 3.75 - 4.00 4.00 - 4.25 4.25 - 4.50 4.50 - 4.75 4.75 - 5.00	4 5 4 2 5 9 4 5 4	4 1 4 2 11 7 5 6 5	1 2 4 8 6 5 11 4 4	0 0 0 0 0 9 8 12 12 22 21 20 15 13 1
	Total	43	46	48	137

Table 4-2. We notice in Table 4-1 that, for each of the three levels, the mode of the total frequency distribution is the highest category, 3.75 through 4.00. If we examine the frequency distributions of separate subtests, however, we will notice that there are some variations among their configurations. Above all, it is noted that Subtests L2, L3 and L4 have different modes from the highest category, i.e., mostly either the category, 3.00 through 3.25, or the category, 3.25 through 3.50. This tendency is also shared by Subtest L1, which has five-alternative multiple-choice test items, as is shown in Table 4-2.

Figure 4-1 presents the scatter diagram of all the fouralternative test items which are included in both Levels 11 and
12, with respect to the values of the estimated Index k*, and
the corresponding scatter diagram for Levels 12 and 13. The
numbers of the shared test items are 264 for Levels 11 and 12,
and 323 for Levels 12 and 13. There is a tendency in both
scatter diagrams that more items are located below the line of
45 degrees, which is drawn in the graphs, and this tendency is
more conspicuous in the first one in which Levels 11 and 12
are compared. This indicates the existence of items whose values
of the estimated Index k* are higher for the lower level and
lower for the higher level. Similar scatter diagrams were
drawn for each of the eleven subtests, including Subtest L1,
and for each of the two pairs of levels, and the resultant twentytwo graphs are presented in Figures 4-2 and 4-3. We can see in

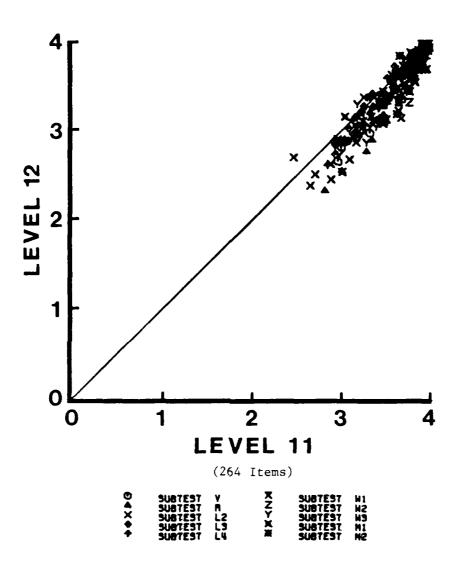


FIGURE 4-1

Comparison of the Values of Index k* for the Four-Alternative Items of the Eleven Subtests Administered to Two Adjacent Levels of Students: Levels 11 and 12.

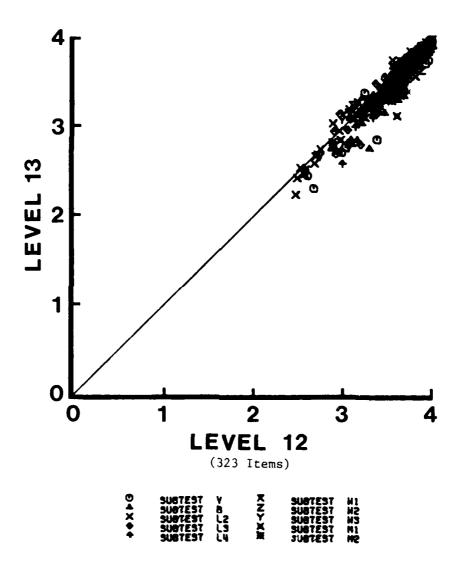
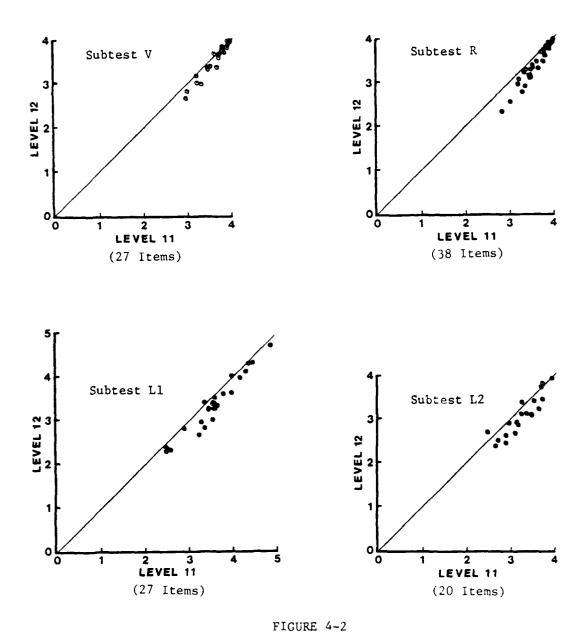
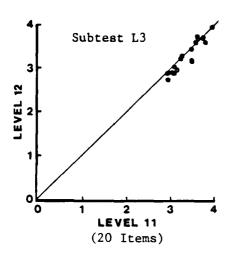
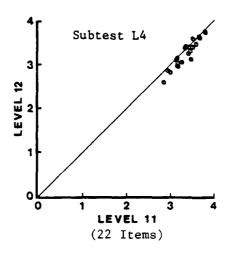


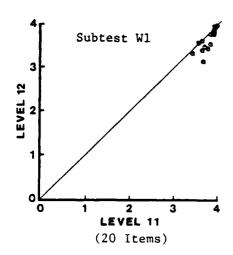
FIGURE 4-1 (Continued): Levels 12 and 13.



Comparison of the Values of Index k* for the Items of Each Subtest Administered to the Students of Both Levels 11 and 12.







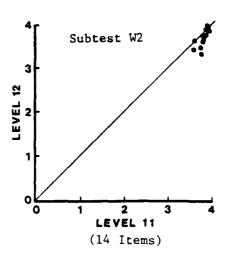
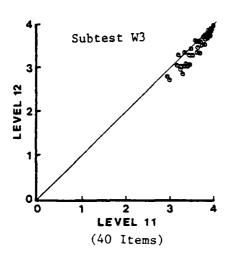
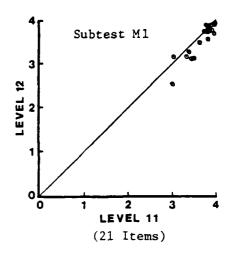


FIGURE 4-2 (Continued): Levels 11 and 12.





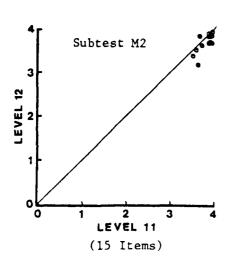
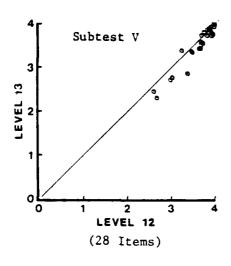
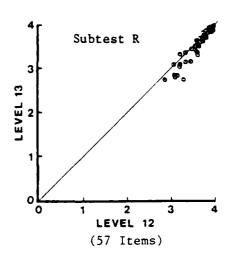
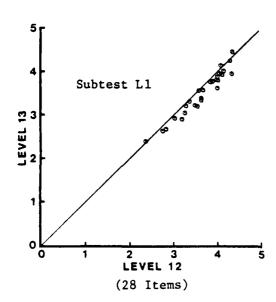


FIGURE 4-2 (Continued): Levels 11 and 12.







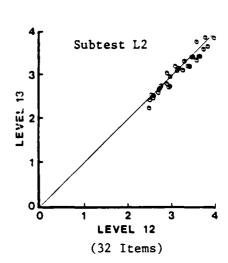
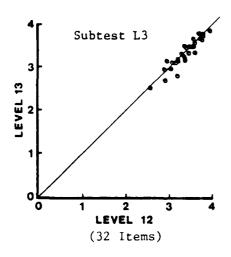
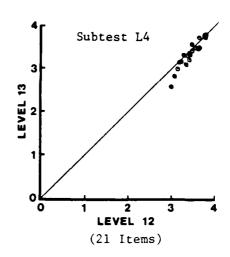
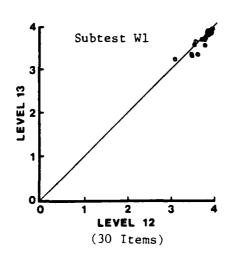


FIGURE 4-3

Comparison of the Values of Index k* for the Items of Each Subtest Administered to the Students of Both Levels 12 and 13.







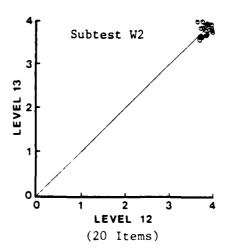
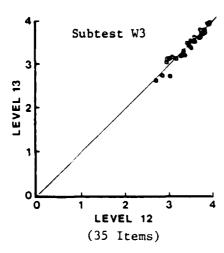
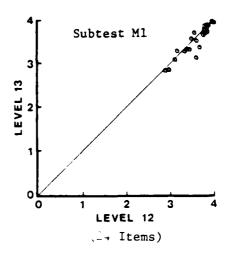


FIGURE 4-3 (Continued): Levels 12 and 13.





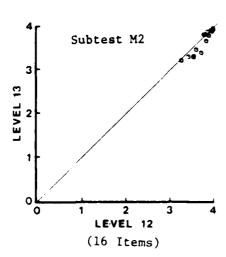


FIGURE 4-3 (Continued): Levels 12 and 13.

this figure that the above tendency exists in all the comparisons.

Table 4-3 presents the identification of each fouralternative test item whose estimated Index k* is 3.9 or greater,
the value of the estimated Index k*, and the probability obtained
as the result of the chi-square test, which was referred to
earlier in this chapter. Note, however, that this probability
was obtained upon the original data, while the estimated Index k*
was obtained upon the revised data. The numbers of fouralternative test items whose estimated Index k* are greater than,
or equal to, 3.9 are 78, 73 and 76 for Levels 11, 12 and 13,
respectively. We can see there is an approximate correspondence
between the probabilities which are greater than, or equal to,
0.1 and the estimated values of Index k* which are 3.9 or greater.

Eighty-two examples of the frequency distribution of the examinees with respect to their choices of an answer out of the four alternatives are presented as Figure 4-4. These test items are the first twenty-four items in each of the three levels in Table 4-3, added by those listed in the same table whose corresponding probabilities obtained as the result of the chi-square test are greater than, or equal to, 0.0005. Thus we have 26, 27 and 29 four-alternative test items for Levels 11, 12 and 13, respectively, to make the total number of illustrated test items 82. In each histogram, also drawn by a dotted line is the estimated proportion, P_R^{\star} , multiplied by the number of examinees who answered the item in one way or another, or the

TABLE 4-3

TABLE 4-3 (Continued): Level 11.

Probability Obtained by the Chi-Square Test (Original Data) against the Uniform Distribution for Items Whose Index k*'s Are 3.9 or Greater (Revised Data). Items Are Arranged in the Descending Order of the Values of Index k*.

Level 11

Subtest 6 Item Number	Value of Index k	Probab111ty
M1-068	3.99933	0.815
W1-021	3.99881	0.824
W3-039	3.99760	0.565
R -078	3.99728	0.889
W1-025	3.99669	0.514
M1-038	3.99635	0.669
M1-037	3.99544	0.121
M1-045	3.99404	0.246
W1-029	3.99401	0.147
V -061	3.99396	0.061
W3-053	3.99237	0.142
M1-031	3.99081	0.445
W3-040	3.99015	0.111
M1-063	3.98452	0.001
R -041	3.98349	0.000
R -047	3.98318	0.063
V -057	3.98067	000.0
R -088	3.97979	0.004
R -073	3.97964	0.002
M1-047	3.97779	000.0
W2-028	3.97392	0.000
M1-055	3.97082	0.000
W2-029	3.96987	0.012
W2-041	3.96974	0.001

25 W3-033 3.96878 26 R -090 3.96878 28 W1-043 3.96448 28 W1-045 3.96329 29 W3-063 3.96061 30 H2-054 3.95814 31 H2-037 3.95814 32 R -057 3.95814 33 W3-028 3.95423 34 W3-028 3.95423 41 W2-043 3.95429 42 W -058 3.9429 43 W3-060 3.94409 44 W2-043 3.9429 45 W -063 3.94799 46 W -063 3.94799 47 W1-046 3.94609 48 R -014 3.93675 49 R -046 3.93657 50 W3-070 3.93457 51 W1-019 3.93369 52 W1-028 3.93369		Subtest & Item Number	Value of Index k*	Probability
R -090 W1-043 W1-043 W1-045 W1-045 W1-045 W1-045 W1-046 W1-046 W1-046 W1-046 W1-046 W1-046 W1-046 W1-046 W1-047 W1-048 W1-047 W1-047 W1-049 W1-047 W1-049 W1-047 W1-049 W1-047 W1-049 W1-047 W1-040 W1	25	W3-033	3.96878	0.000
M1-043 M2-045 M3-063 M2-042 M1-042 M1-042 M1-043 M1-044 M1-046 M1-046 M1-046 M1-047 M1-047 M1-047 M1-019 M1-028 M1-029	26	R -090	3.96586	0.000
H2-045 H2-053 H2-054 H2-054 H2-054 H2-054 H2-042 H2-042 H2-042 H2-043 H2-043 H2-043 H3-060 H3-070 H3-070 H3-070 H3-070 H3-070	27	W1-043	3.96448	0.000
H2-063 H2-037 H2-037 H2-037 H2-042 H3-028 H3-028 H3-028 H3-020 H3-050 H3-050 H3-050 H3-060 H2-043 H2-043 H2-043 H2-043 H2-043 H2-043 H2-043 H2-043 H2-043 H3-070 H3-070 H3-070 H3-070 H3-070 H3-070 H3-070 H3-070	28	M1-045	3.96329	000.0
H2-054 H2-037 R -057 H2-042 H2-042 H3-028 H3-028 H3-049 H2-043 H3-049 H2-043 H3-040 H3	29	W3-063	3.96061	000.0
H2-037 R -057 H2-042 12-058 H3-028 H3-020 H3-050 H2-043 H2-045 H2-043 H2-045 H2-043 H2-045 H2-043 H3-050 H3-070 H3-070 H3-070 H3-070 H3-070	30	M2-054	3.95937	000.0
R - 057 H2-042 12-058 H3-028 H3-028 H3-029 H3-050 H3-050 H2-043 H3-060 H3-070 H3-070 H3-070 H3-070 H3-070 H3-070	31	H2-037	3.95814	000.0
H2-042 12-058 13-028 13-028 13-028 13-028 13-020 13	32		3.95716	000.0
12-058 13-028 13-028 13-029 13-020 13	33	M2-042	3.95493	0000
W3-028 L3-049 W1-020 W1-020 W1-026 W2-045 W2-045 W2-045 W2-043 W1-060 W1-047 W1-047 W1-047 W1-047 W1-028 W1-028 W1-028 W1-028 W1-028 W1-047 W1-047 W1-047 W1-040 W1-040	34	12-058	3.95473	000.0
L3-049 W1-020 W1-020 W1-020 W1-046 W2-043 W2-043 W1-045 W1-043 W1-047 W1-047 W1-019 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028	35	W3-028	3.95427	00.000
W1-020 W1-046 W1-046 W1-043 W2-043 W1-060 W1-043 W1-047 W1-047 W1-047 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028	36	L3-049	3.95320	0.000
W1-046 W1-046 W2-043 W2-045 W3-060 W2-043 W1-047 W1-047 W1-047 W1-019 W1-019 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028	37	W1-020	3.95258	0.036
W1-046 M2-043 W2-045 W2-043 W2-043 W2-043 W1-047 W1-047 W1-047 W1-019 W1-019 W1-028 W1-028 W1-049 W1-049 W1-049 W1-049 W1-049 W1-049 W1-049	38	W3-050	3.95042	000.0
H2-043 W2-045 V -058 W3-060 W2-043 V -063 V -048 W1-047 R -031 R -042 W1-019 W1-028 W1-028 W1-028 W1-028 W1-028 W1-028 W1-049	39	W1-046	3.94947	0000
W2-045 V -058 W3-060 W2-043 W1-047 W1-047 W3-070 W3-070 W1-019 W1-019 W1-028 W1-028 W1-028 W1-028 W1-049 W1-028	0,7	M2-043	3.94942	00000
V -058 W3-060 W2-043 V -063 V -048 W1-047 R -031 R -042 W3-070 W1-019 W1-028 V -049	41	W2-045	3.94929	000.0
W3-060 W2-043 W2-043 W1-047 W1-047 W3-070 W3-070 W1-019 W1-028 W1-049	42	v -058	3.94799	0000
W2-043 3. V -063 3. V -063 3. V -063 3. V -048 3. V -048 3. V -047 3. V -042 3. V -049	43	M3-060	3.94409	000.0
V -063 3. W1-047 3. R -031 3. R -042 3. W3-070 W1-019 3. V -049 3.	77	W2-043	3.94234	0000
V -048 W1-047 3. R -031 3. R -042 3. W3-070 3. W1-019 3. V -049	4.5		3.94127	0000
W1-047 3. R -031 3. R -042 3. W3-070 3. W1-019 3. V -049 3.	97		3.93675	000.0
R -031 3. R -042 3. W3-070 3. W1-019 3. V -049 3.	47	W1-047	3.93600	000.0
R -042 3. W3-070 3. W1-019 3. V -049 3.	84		3.93572	0000
W1-019 3. W1-028 3. V -049 3.	67		3.93549	0.081
W1-019 3.9 W1-028 3.9 V -049 3.9	20	W3-070	3.93457	000.0
W1-028 3.9 V -049 3.9	51	W1-019	3.93369	000.0
V -049 3.9	52	W1-028	3.93036	000.0
	53		3.92777	0.000

0.018

0.009 0.005 0.000 0.004 0.001 0.000 0.001 0.009 0.000

0.000 0.001

TABLE 4-3 (Continued): Level 11,

Subtest & Item Number

R -080

W1-037

V -052

M3-069 W3-030 W3-078 W1-040 M1-069 W3-031 M2-048 W2-044

V -054

Probability

TABLE 4-3 (Continued): Level 12.

0.383 0.048 0.171 0.261 0.022 0.007

Value of Index k*	3.99933	3.99512	3.99508	3.99508	3.99274	3.99234	3.99209	3.99194	3.99092	3.98762	3.98608	3.98370	3.98333	3.98260	3.98129	3.97931	3.97886	3.97850	3.97762	3.97637	3.97371	3.97358	3.97239	3.96896	
Subtest & Item Number	M1-095	в -078	W1-029	V -079	W3-053	V -057	H1-074	W2-044	W1-053	M2-059	R -131	670- A	W2-048	V -083	W2-059	V -052	M2-060	R -127	R -088	W3-083	R -105	W2-047	W1-028	W3-098	
	7	7	~	4	Ş	9	^	80	6	01	11	12	13	14	115	16	17	18	19	20	21	22	23	77	
Probability	0.000	0.000	0.000	0.00	0.000	0.000	0.000	0.000	0.000	0.000	0000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0000	0.000	0.000	0000	0.000

W1-031

R -096 R -097 W1-038 W1-035 R -030 R -032 W3-044 W1-018 W3-047

M1-061

TABLE 4-3 (Continued): Level 12.

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Probability	00000	0.000	0.000	0.000	0.000	0.001	0.000	0.001	0.000	0.000	0.000	0.000	0.000	0.000	0.016	0.000	0.000	000.0	0.000	0.000	0.000	0.000	000.0	000.0	0.000	000.0	0.000	0.000	0.000
Value of Index k	3.96849	3.96644	3.96614	3.96585	3.96546	3.96525	3.96483	3.96465	3.96262	3.96052	3.95866	3.95746	3.95658	3.95600	3.95557	3.95499	3.95318	3.95215	3.95203	3.95174	3.94713	3.94622	3.93945	3.93684	3.93447	3.92976	3.92899	3.92824	3.92556
Subtest & Item Number	R -116	M2-054	v -061	M1-084	R -128	W2-045	R -103	W3-050	W1-043	M2-066	W1-063	W1-054	L3-049	V -076	M1-068	W1-045	12-058	M1-088	W1-067	W1-062	W1-057	W3-093	R -134	W1-046	M2-069	R -133	R -123	R -073	42-046
	25	56	27	28	29	99	31	32	33	34	35	%	37	38	39	0,7	1,7	42	63	77	4.5	9,7	47	8,7	67	20	5.1	52	53

TABLE 4-3 (Continued): Level 12.

	Subtest & Item Number	Value of Index k	Probability
54	R -132	3.92516	0.000
55	W3-070	3.92434	000.0
56	V -058	3.92251	000.0
57	R -124	3.92229	0.000
28	W2-049	3.92155	000.0
59	₩1-066	3.92136	000.0
09	M1-063	3.91931	0.000
61	R -122	3.91425	0.000
62	v -063	3.91414	0.000
63	W1-061	3.91392	0.000
79	W2-060	3.91359	0.000
65	В -092	3.91305	0.000
99	W3-063	3.90916	0.000
19	R -129	3.90905	0.000
89	M1-061	3.90706	0.000
69	R -090	3.90414	0.000
70	M1-057	3.90347	0.000
11	R -118	3.90282	0.000
72	W1-051	3.90055	0.000
73	W2-037	3.90046	0.000

TABLE 4-3 (Continued): Level 13.

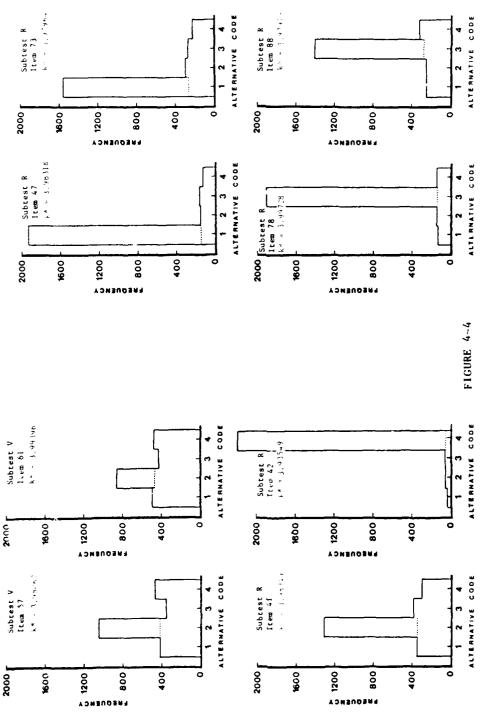
Level 13.	Probability	0.851	0.915	0.403	0.280	0.249	0.186	0.050	0.122	0.066	0.096	0.021	0.063	0.098	0.134	0.000	000.0	0.000	000.0	0.000	0.000	000.0	0.002	0.000	0.089
	Value of Index k*	3.99961	3.99953	3.99820	3.99691	3.99662	3.99598	3.99525	3.99401	3.99350	3.99311	3.99297	3.99110	3.99105	3.98949	3.98035	3.97878	3.97836	3.97714	3.97500	3.97449	3.97339	3.97322	3.97137	3.97129
TABLE 4-3 (continued):	Subtest & Item Number	V -093	M2-044	R -140	M1-107	R -155	W3-083	M2-075	V -061	W1-078	W2-048	R -142	v -079	R -131	M1-105	M2-060	W2-066	M2-082	W2-067	W1-075	W1-053	R -150	R -103	M2-078	w3-070
	 																								

Value of Index k Probability 3.96820 0.000 3.96820 0.000 3.96594 0.000 3.96599 0.000 3.96428 0.000 3.96428 0.000 3.96436 0.000 3.9548 0.000 3.95847 0.000 3.95848 0.000 3.95479 0.000 3.95472 0.000 3.95396 0.000 3.95396 0.000 3.95147 0.000 3.95147 0.000 3.95103 0.000 3.95033 0.000 3.95031 0.000 3.95031 0.000
3.94961 0.000 3.94521 0.000

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-

TABLE 4-3 (Continued): Level 13.

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Probability	00.00	00000	0.000	0.000	0.000	0.000	0.000	0.000	0000	000.0	000.0	000.0	0.000	000.0	000.0	000.0	000.0	0.000	000.0	000.0	000.0	000.0	0.000
Value of Index k*	3.94079	3.94043	3.94017	3.93921	3.93713	3.93531	3.93418	3.93004	3.92916	3.92723	3.92541	3.92402	3.91678	3.91674	3.91623	3.91549	3.91517	3.90854	3.90724	3.90657	3.90498	3,90441	3.90411
Subtest & Item Number	W2-063	R -100	v -092	W2-064	V -078	W1-063	M2-069	M2-112	W2-047	W1-060	W2-041	W2-059	W1-068	R -129	R -118	R -133	990-TA	141 - (164	W1-046	W2-068	M2-066	990- A	R -130
	24	55	95	5.7	5.8	65	09	ij	62	63	79	6.5	99	67	89	69	70	71	72	7.3	7.4	7.5	7.6



Frequency Distribution of the Examinees of Each of the Three Levels with Respect to Their Responses to Each Test Item Whose Index k^* Is 3.9 or Greater, with the Estimated Proportion of the Examinees Guessing Correctly (Dotted Line).

Level 11

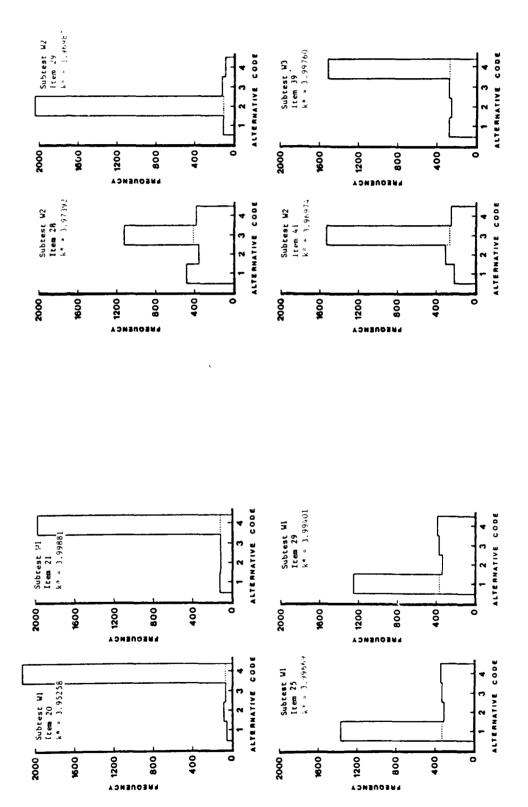


FIGURE 4-4 (Continued): Level 11.

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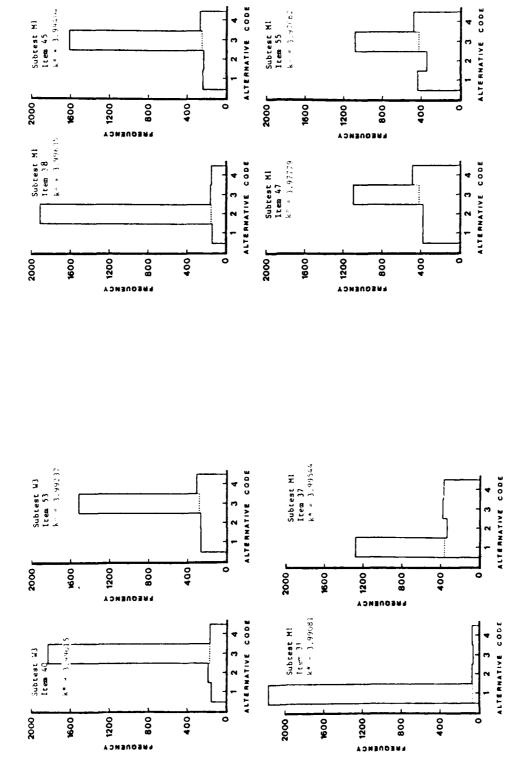
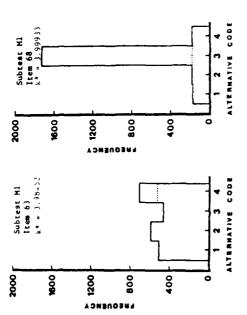


FIGURE 4-4 (Continued): Level 11.





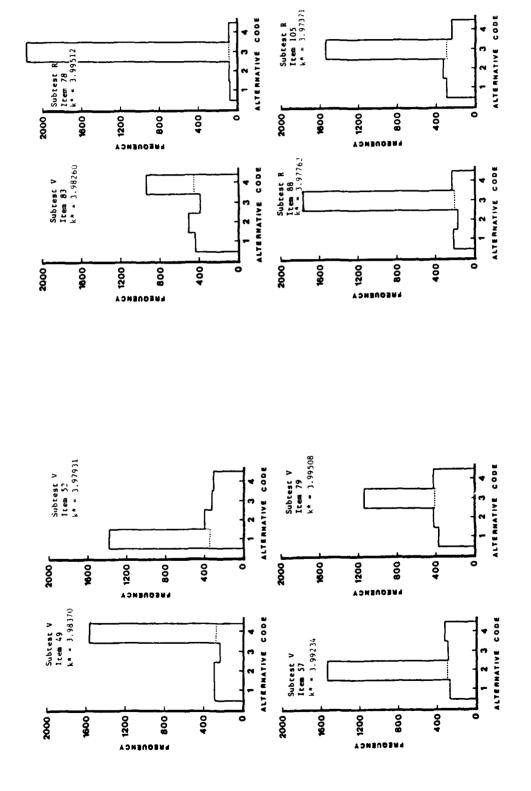


FIGURE 4-4 (Continued): Level 12.

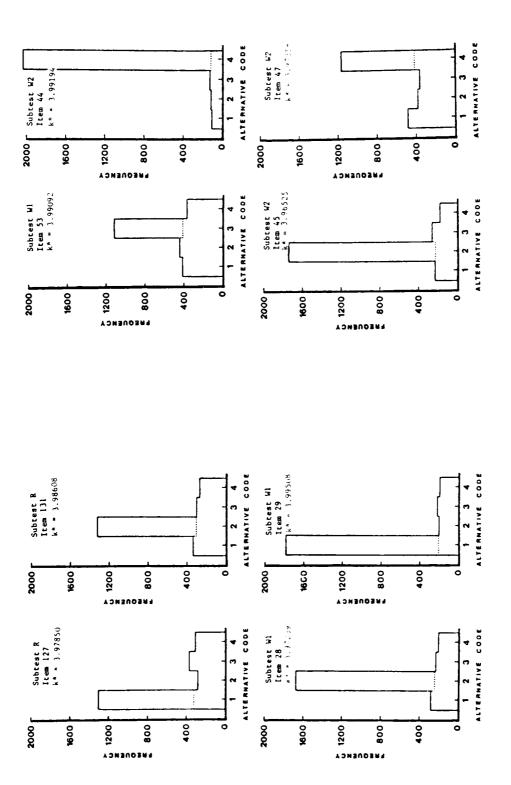


FIGURE 4-4 (Continued): Level 12.

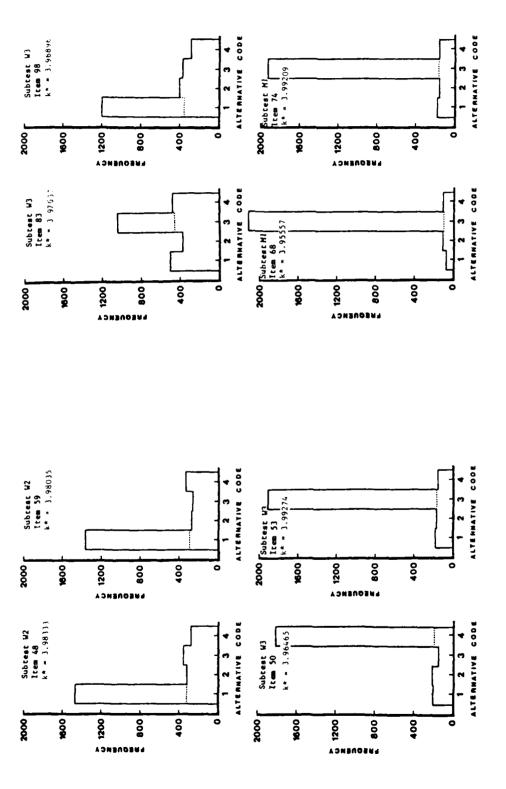


FIGURE 4-4 (Continued): Level 12.

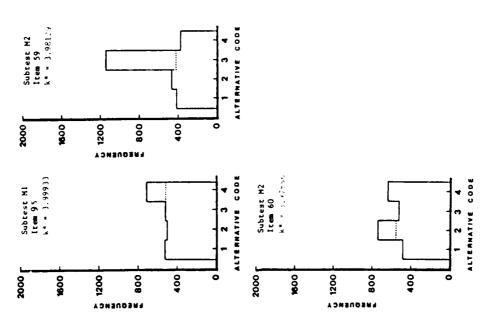


FIGURE 4-4 (Continued): Level 12.

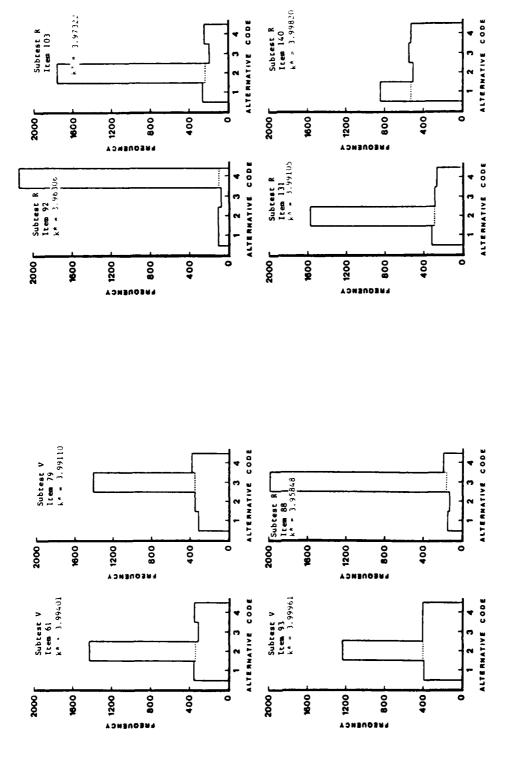
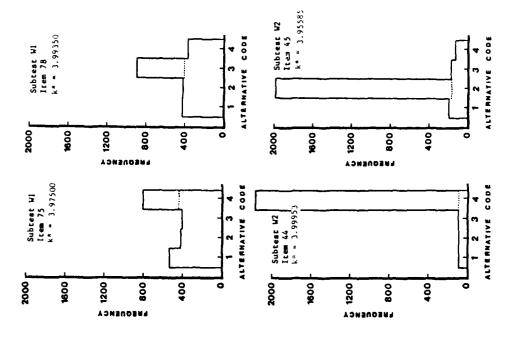


FIGURE 4-4 (Continued): Level 13.



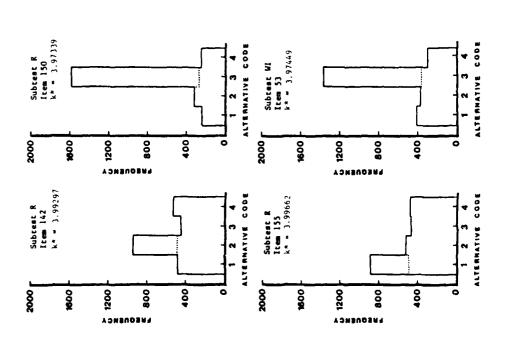
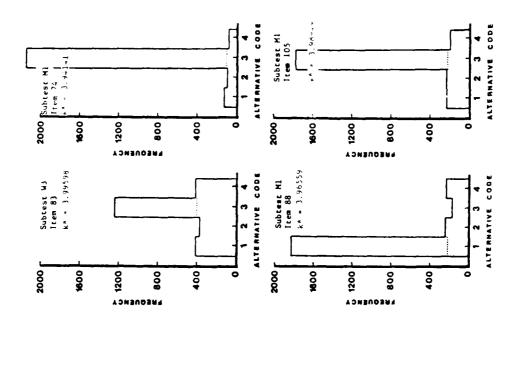


FIGURE 4-4 (Continued): Level 13.



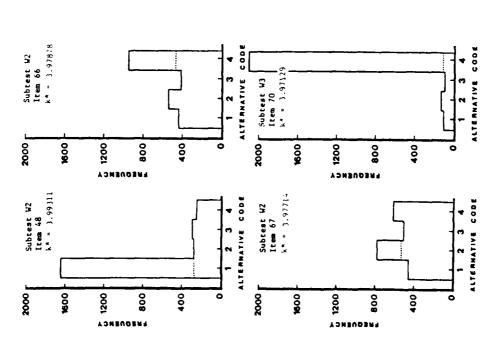
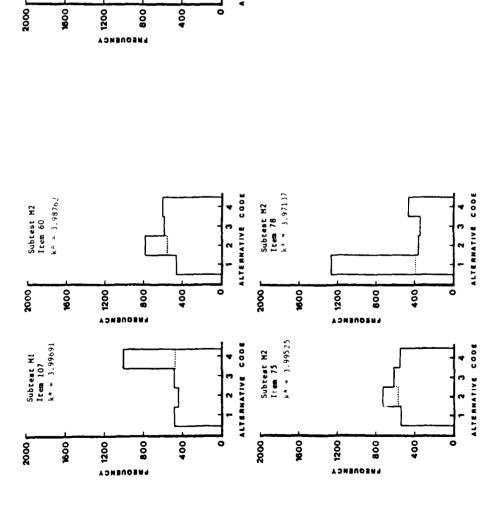


FIGURE 4-4 (Continued): Level 13.

Subtest M2 Item 82 k* = 3.97836



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FIGURE 4-4 (Continued): Level 13.

total number of examinees subtracted by the number of those who did not answer the item at all. We can see in this figure that most of these histograms are close to rectangles, if we replace the frequency for the correct answer by the height indicated by the dotted line in each histogram. It is especially true if we focus our attention to those items whose estimates of Index k* are greater than 3.99, i.e., V-61, R-78, W1-21, W1-25, W1-29, W3-39, W3-40, W3-5; M1-31, M1-37, M1-38, M1-45 and M1-68of Level 11, V-57, V-79, R-78, W1-29, W1-53, W2-44, W3-53, M1-74 and M1-95 of Level 12, and V-61, V-79, V-93, R-131, R-140, R-142, R-155, W1-78, W2-44, W2-48, W3-83, M1-107 and M2-75 of Level 13. On the other hand, if we shift our attention to those whose values of the estimated Index k* are relatively low, like W2-28 of Level 11 (k*=3.97392), W3-83 of Level 12 (k*=3.97637) and M2-78 (k*=3.97137) of Level 13, we shall find some substantial deviations from rectangularity. For the group as a whole, however, these histograms are close enough to rectangles, as we expect from the high values of estimated Index k*.

It should be noted that among these eighty-two four-alternative test items we find no items from Subtest L2, L3 or L4. These subtests are all developed for measuring language skills. In fact, in the total group of 227 test items, whose values of the estimated Index k* are greater than 3.9, there are only four test items from these three subtests, i.e., L2-58 (k*=3.95473) and L3-49 (**=3.7525) at large 11, and

L3-49 (k*=3.95658) and L2-58 (k*=3.95318) of Level 12, which are actually two items shared by both Levels 11 and 12. This result is no surprise, considering the fact that for these subtests the modes of the estimated Index k* are much lower than those of the other subtests, as we have seen earlier in Table 4-1. A close examination of the contents of the test items of these four subtests, including Subtest L1, and their results of analysis reveals the following facts.

- (1) All the questions in these four language skill subtests are in the form of having the examinee find mistakes in spelling, capitalization, punctuation and usage, respectively.
- (2) Unlike the test items in the other seven subtests, these items have "No mistakes" as the last alternative, and for most items this alternative has a high frequency, even when it is a wrong answer.

From these facts and the above results, it is obvious that Equivalent Distractor Model is not suitable for the items of the four subtests of language skills, including Subtest L1, which consists of five-alternative test items. For these items, Informative Distractor Model may be more appropriate.

It should also be noted that, except for a few items like V-61 and M1-63 of Level 11, M1-95 and M2-60 of Level 12, and R-140, R-155, W1-75, W1-78, W2-67, M2-60, M2-75 and M2-82

of Level 13, these items presented in Figure 4-4 are relatively easy items for the respective groups of examinees. In such a case, we cannot expect this type of frequency distribution to reveal the information each alternative has, even if the item belongs to Informative Distractor Model. If we exclude all the four-alternative test items which belong to one of the three subtests of language skills, there are 76 items out of 306 for Level 11, 71 items out of 325 for Level 12, and 76 out of 334 for Level 13, whose estimates of Index k* are 3.9 or greater. These numbers are interpreted as 24.8, 21.8 and 22.8 in percentage for Levels 11, 12 and 13, respectively, for the remaining seven subtests. They are by no means large numbers, however.

estimated Index k*, and the probability obtained by the chi-square test for each of the four-alternative test items whose estimates of Index k* are the lowest, i.e., 3.0 or less. These items number 134, of which 42 are of Level 11, 49 are of Level 12, and the remaining 43 are of Level 13. We can see in this table that all the probabilities, which were obtained upon the original data, are less than 0.0005. It is noted that there are many items from Subtest L2, L3. or L4, unlike those for the highest Index k* group, which are shown in Table 4-3. In fact, there are as many as 27 such items for Level 11, 31 for Level 12 and 26 for Level 13. Out of these test items, 7, 11 and 9 items

TABLE 4-4

Probability Obtained by the Chi-Square Test against the Uniform Distribution for Item Alternatives Whose Index k* Values Are 3.0 or Leas.
Items Are Arranged in the Descending Order of the Values of Index k*.

Level 11

																						_		
Probability	0.000	0.000	000.0	0.000	0.000	00000	0000	0.000	0.000	0.000	0.000	000.0	0.000	000.0	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Value of Index k*	2.98574	2.98446	2.98120	2.98077	2.96675	2.96582	2.96109	2.95480	2.95337	2.94519	2.94137	2.93717	2.92684	2.91176	2.90829	2.90430	2.90014	2.88903	2.87004	2.83820	2.83691	2.83280	2.83258	2.78061
Subtest & Item Number	090- A	L3-038	L2-053	W2-022	W2-026	L4-038	13-053	L3-041	W3-043	L3-033	L3-022	L3-052	V -027	L4-024	L2-046	L2~048	L2-029	L2-027	L4-041	R -063	L2-023	R -037	L3-036	V -034
	1	2	~	4	\$	٠	7		5	10	11	12	13	14	1.5	16	17	18	19	20	21	22	23	24

TABLE 4-4 (Continued): Level 11.

	Subtest & Item Number	Value of Index k*	Probability
25	L3-020	2.77848	000.0
26	12-054	2.72988	0.000
27	M1-032	2.72056	000.0
28	L2-035	2.71827	000.0
29	12-034	2.70963	0.000
8	L2-040	2.67619	0.000
31	L3-025	2.65321	0.000
32	L2-025	2.63903	0.000
33	12-028	2.63035	0.000
34	V -024	2.62267	0.000
35	R -045	2.61763	0.000
36	R -050	2.58812	0.000
37	12-019	2.54288	0.000
38	W1-027	2.52326	000.0
39	L2-055	2.49083	0.000
07	17-054	2.47698	000.0
41	R -034	2.45910	000.0
42	W2-023	2.31156	000.0

0.000

0.000

TABLE 4-4 (Continued): Level 12.

Probability

TABLE 4-4 (Continued): Level 12.

0.000

Value of Index k*	2.84096	2.83073	2.78972	2.75618	2.75246	2.75097	2.72287	2.70300	2.70073	2.68977	2.68016	2.64858	2.62909	2.61062	2.58695	2.58167	2.57592	2.56920	2.56018	2.53272	2.49724	2.47992	2.47486	2.40572	2.34781
Subtest 6 Item Number	v -053	W3-043	R -065	L3-041	12-011	W3-068	1.2-055	17-080	12-041	1.2-078	090- A	L2-048	14-041	V -084	17-070	L3-072	12-063	R -066	M1-054	12-054	12-079	17-067	17-046	L2-040	R -063
	25	26	27	28	29	28	31	32	33	34	35	36	37	38	33	70	41	42	43	7,7	4.5	9,7	47	87	49
Probability	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0000	0.000	
Value of Index k*	2.99099	2,98965	2.98524	2,98363	2.97425	2.96648	2.96261	2.95951	2.95558	2.92803	2.92549	2.92517	2.91927	2,90879	2.89935	2,89917	2.89777	2.89085	2.89001	2.88979	2.88760	2.88622	2.87424	2.85821	
Subtest & Item Number	W3-072	V -071	L4-034	L3-048	M1-085	R -067	L3-069	12-068	12-056	12-053	L3-046	L3-053	R -071	L3-040	L3-042	13-052	L4-038	W3-084	W3-045	12-049	M1-091	12-061	R -106	1.4-043	

0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000

000.0

000.0

TABLE 4-4 (Continued): Level 13.

Probability	000.0	000.0	0.000	000.0	000.0	000.0	000.0	00.00	00000	00.00	0000	0.000	000.0	00.00	0000	0.000	000.0	0.000	0.000	000.0	00.00	00.00	00.00	00.00
Value of Index k*	2.98990	2.97913	2.97289	2.92099	2.88186	2.87719	2.87223	2.87204	2.87129	2.84972	2.84871	2.83841	2.82643	2.82386	2.82275	2.81845	2.79924	2.79228	2.77642	2.77484	2.77466	2.77141	2.76178	2.75648
Subtest & Item Number	L3-054	17-068	13-052	14-071	V -059	R -109	M1-085	12-085	R -082	M1-091	14-050	12-081	L3-086	R -083	13-055	12-061	W3-084	790- A	W3-067	R -087	R -106	1.2-077	12-056	17-088
	1	7	٣	4	'n	9	^	60	6	10	17	12	13	14	15	16	17	18	19	20	21	22	23	24

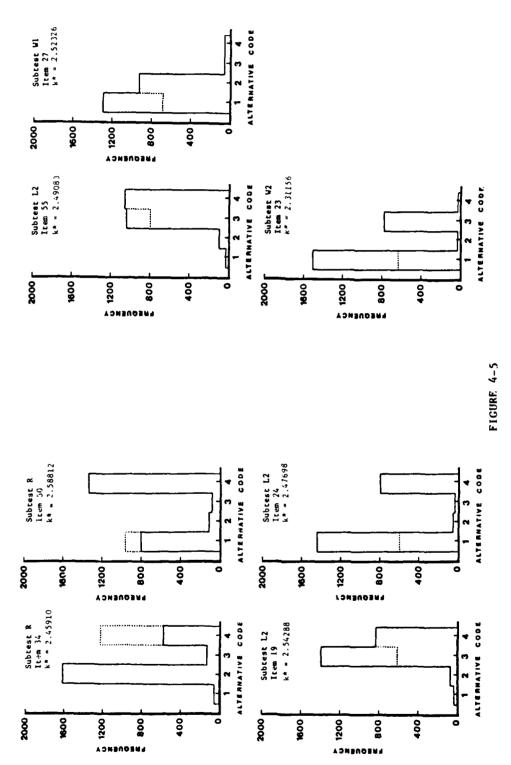
1
Value of
Subtest &
1

TABLE 4-4 (Continued): Level 13.

	Subtest & Item Number	Value of Index k*	Probability
25	12-053	2.73676	000.0
26	v -071	2.73664	000.0
27	L3-053	2.71696	0000
28	12-055	2.71265	000.0
29	960- A	2.69742	00000
30	17-080	2.69277	0.000
31	W3-068	2.68418	0000
32	17-090	2.63891	0.000
33	1.2-078	2.61208	000.0
75	770-77	2.60493	000.0
33	12-054	2.56223	0.000
36	12-070	2.54443	0.000
37	L3-072	2.54169	0.000
88	660- A	2.50520	0.000
39	1.2-063	2.49213	0000
3	V -084	2.47745	0.000
7,	12-079	2.44666	000.0
42	090- A	2.32986	0.000
43	17-067	2.25958	000.0

have the values of the estimated Index k* less than, or equal to, 2.6, respectively, for Levels 11, 12 and 13. Histograms similar in nature as those in Figure 4-4 are drawn for the frequency distributions of these 27 four-alternative test items, and are presented as Figure 4-5. We can see that these histograms, whose frequencies for the correct answer are replaced by the dotted lines, are by no means close to rectangles. There is no reason to accept Equivalent Distractor Models for these test items. Among them, there are such items as R-34 and R-50 for Level 11, L2-79 and L3-72 for Level 12, and L2-79 and L3-72 for Level 13 whose values of P* multiplied by the respective numbers of examinees who answered in one way or another exceed the corresponding frequencies for the correct answer; in these cases, the use of Index k* itself is meaningless.

Figure 4-6 presents the frequency distributions of six four-alternative test items for each of the three levels, which were arbitrarily selected from those whose values of the estimated Index k* are greater than, or equal to, 3.0 and less than 3.6. They are samples from the four-alternative test items with intermediate values of the estimated Index k*. We can see that these histograms, in which the frequencies for the correct answer are replaced by the dotted lines, are also far from rectangles. They also include one item, i.e., L4-74 for Level 13, for which the dotted line exceeds the frequency



Frequency Distribution of the Examinees of Each of the Three Levels with Respect to Their Responses to Each Test Item Whose Index k* is 2.6 or Less, with the Estimated Proportion of the Examinees Guessing Correctly (Dotted Line).

Level 11

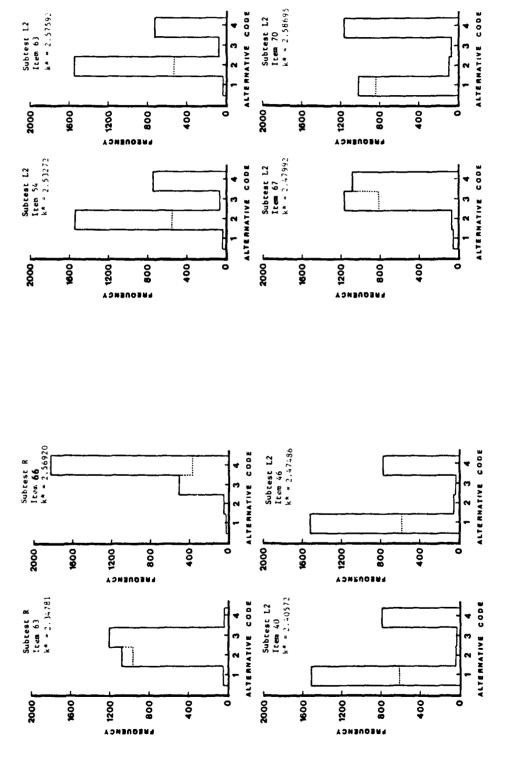
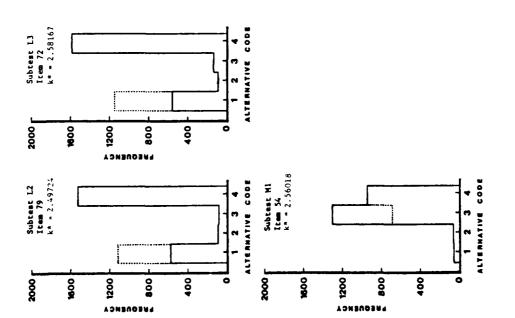


FIGURE 4-5 (Continued): Level 12.





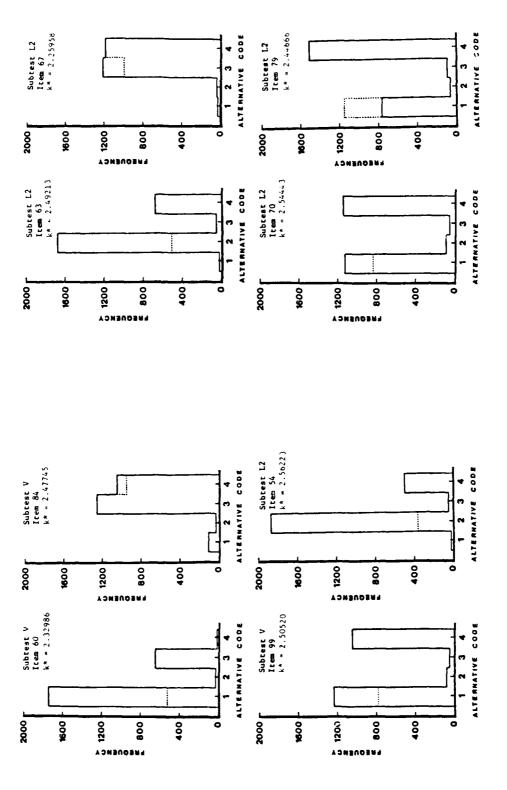
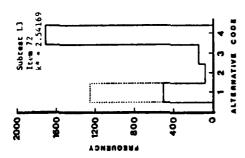


FIGURE 4-5 (Continued): Level 13.





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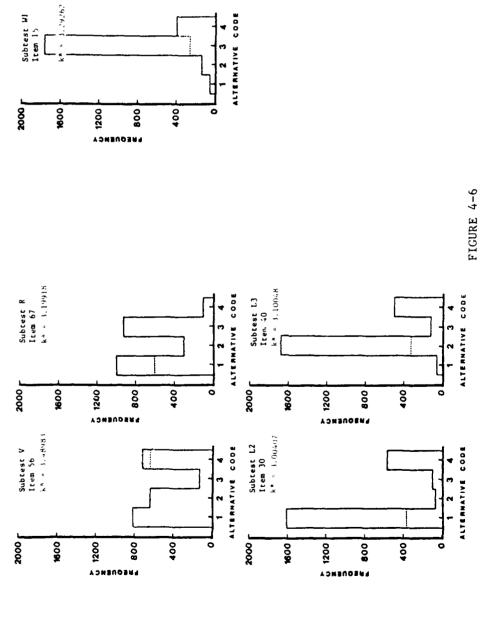
Subtest W3
Item 81
k* = 3.40849

2000 g

1000

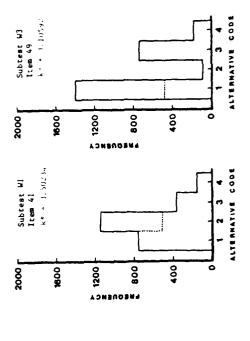
##EQUENCY

400



Frequency Distribution of the Examinees of Each of the Three Levels with Respect to Their Responses to Each Test Item Whose Index k* Is Greater than, or Equal to, 3.0 and Less than 3.6 , with the Estimated Proportion of the Examinees Guessing Correctly (Dotted Line).

Level II



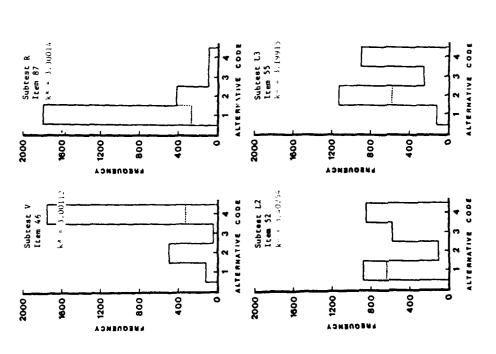


FIGURE 4-6 (Continued): Level 12.

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7

Subtest M1 Item 89

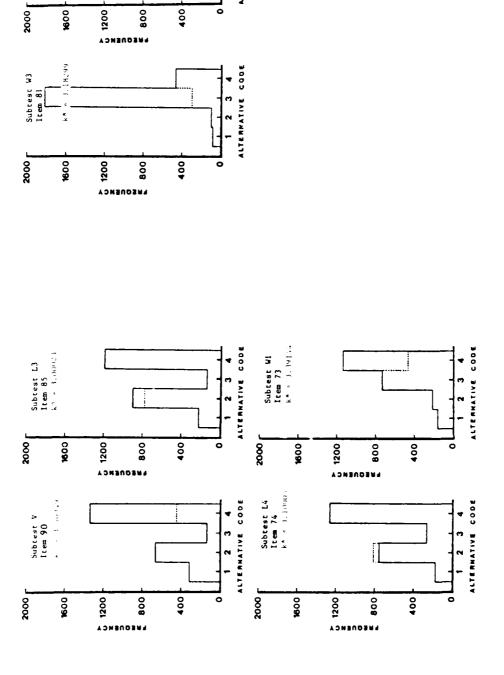
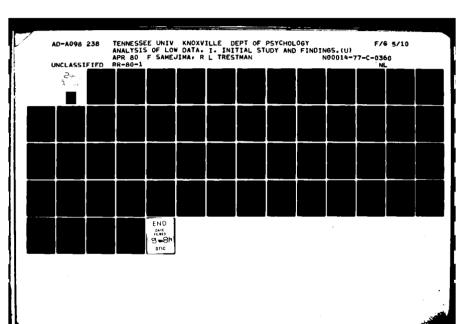
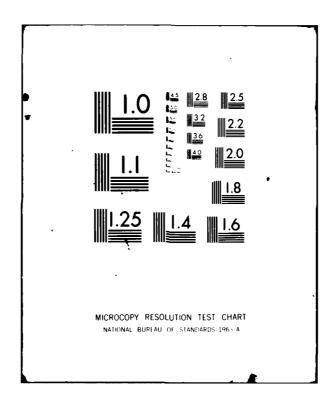


FIGURE 4-6 (Continued): Level 13.

for the correct answer.

As an additional information, Figure 4-7 presents similar histograms for W1-13 and R-26 of lever 11, since their in malifities obtained by the chi-square test, which was conducted upon the original data, turned out to be 0.005 and 0.001, respectively, although their values of estimated index k% are less than 1.4, i.e., 3.87789 for W1-15 and 3.83109 for R-26.





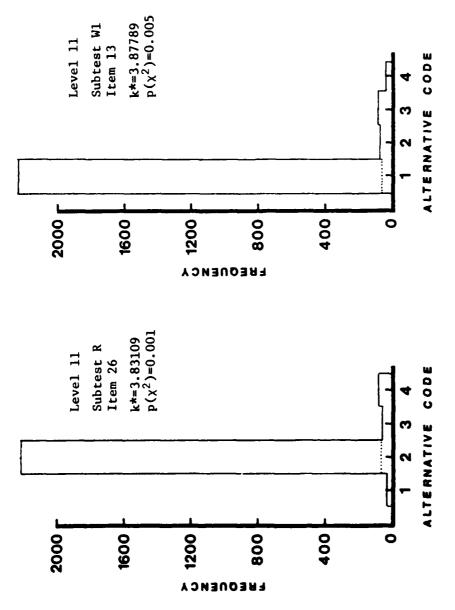


FIGURE 4-7

Frequency Distribution of Examinees with Respect to Their Responses to Each Test Item Whose Index k* Is Less Than 3.9 (Revised Data) and Whose Probability Obtained by the Chi-Square Test (Original Data) is 0.001 or Greater, with the Estimated Proportion of the Examinees Guessing Correctly (Dotted Line).

V Comparisons of the Results of Test Items Which Were Administered to Levels 11, 12, and 13

There are certain test items which are included in all the three levels. They are items 58 through 66 of Subtest V, 80 through 98 of Subtest R, 58 through 66 of Subtest L1, 49 through 58 of Subtest L2, 49 through 58 of Subtest L3, 44 through 54 of Subtest L4, 38 through 47 of Subtest W1, 41 through 46 of Subtest W2, and 67 through 82 of Subtest W3. Thus there are nine items of Subtest V, which were administered to all the three groups of students, nineteen of Subtest R, nine of Subtest L1, ten of Subtest L2, ten of Subtest L3, eleven of Subtest L4, ten of Subtest W1, six of Subtest W2 and sixteen of Subtest W3, which make the total number of test items shared by all the three levels of test 100. There is no item which is included in all three levels for Subtests M1 and M2.

It is evident that, for the behavior of the test item to follow Equivalent Distractor Model, not only the value of estimated Index k* should be close to m for one level of examinees but also for all three levels. For this reason, it will be worthwhile to compare the results across the three levels for these one hundred test items which are included in all the three levels of test.

Tables 5-1 and 5-2 present the three values of estimated

Index k*, which were rounded to the second digit after the decimal point, for each of the ninety-one four-alternative test items, and for each of the nine five-alternative test items, respectively.

We can see that only seven four-alternative test items, i.e., V-61,

TABLE 5-1

Three Different Values of Index k* for Each of the Four Alternative Items That Were Administered to All Three Levels, with Their Respective Item Numbers.

Subtest ₹

I	tem Numbers		Level	
	(11)(12)(13)	11	12	13
58	35 19 1	3.95	3.92	3.73
59	36 20 2	3.69	3.38	2.88
60	37 21 3	2.99	2.68	2.33
61	38 22 4	3.99	3.97	3.99
62	39 23 5	3.74	3.68	3.53
63	40 24 6	3.94	3.91	3.82
64	41 25 7	3.24	3.03	2.79
65	42 26 8	3.86	3.72	3.57
66	43 27 9	3.82	3.86	3.90

TABLE 5-1 (Continued): Subtest R.

I	tem Numbers		Level	
	(11)(12)(13)	11	12	13
80	56 20 1	3.93	3.86	3.83
81	57 21 2	3.53	3.35	3.17
82	58 22 3	3.44	3.16	2.87
83	59 23 4	3.46	3.10	2.82
84	60 24 5	3.80	3.61	3.39
85	61 25 6	3.62	3.48	3.46
86	62 26 7	3.22	3.08	3.12
87	63 27 8	3.37	3.30	2.77
88	64 28 9	3.98	3.98	3.96
89	65 29 10	3.80	3.74	3.63
90	66 30 11	3.97	3.90	3.86
91	67 31 12	3.80	3.65	3.53
92	68 32 13	3.88	3.91	3.96
93	69 33 14	3.79	3.74	3.67
94	70 34 15	3.82	3.61	3.33
95	71 35 16	3.89	3.82	3.72
96	72 36 17	3.91	3.77	3.72
97	73 37 18	3.91	3.81	3.74
98	74 38 19	3.78	3.75	3.83

TABLE 5-1 (Continued): Subtest L2.

It	em Numbers		Level	
	(11X12X13)	11	12	13
49	31 11 1	3.19	2.89	3.06
50	32 12 2	3.71	3.75	3.87
51	33 13 3	3.66	3.25	3.13
52	34 14 4	3.27	3.40	3.22
53	35 15 5	2.98	2.93	2.74
54	36 16 6	2.73	2.53	, 2.56
55	37 17 7	2.49	2.72	2.71
56	38 18 8	3.15	2.96	2.76
57	39 19 9	3.38	3.14	3.17
58	40 20 10	3.95	3.95	3.86

TABLE 5-1 (Continued): Subtest L3.

1	tem Numbers		Level							
	(11X12X13)	11	12	13						
49	31 11 1	3.95	3.96	3.86						
50	32 12 2	3.75	3.72	3.80						
51	33 13 3	3.80	3.61	3.36						
52	34 14 4	2.94	2.90	2.97						
53	35 15 5	2.96	2.93	2.72						
54	36 16 6	3.10	3.05	2.99						
55	37 17 7	3.49	3.20	2.82						
56	38 18 8	3.62	3.70	3.80						
7	39 19 9	3.58	3.61	3.56						
8	40 20 10	3.28	3.30	3.33						

TABLE 5-1 (Continued): Subtest L4.

I:	tem Numbers		Level	
	(11)(12)(13)	11	12	13
44	22 12 1	3.18	3.00	2.60
45	23 13 2	3.47	3.33	3.12
46	24 14 3	3.82	3.76	3.79
47	25 15 4	3.53	3.61	3.48
48	26 16 5	3.18	3.18	3.17
49	27 17 6	3.44	3.28	3.34
50	28 18 7	3.19	3.07	2.85
51	29 19 8	3.52	3.42	3.35
52	30 20 9	3.81	3.78	3.75
53	31 21 10	3.69	3.63	3.75
54	32 22 11	3.35	3.41	3.39

TABLE 5-1 (Continued): Subtest W1.

	tem Numbers		Level	
	(11)(12)(13)	11	12	13
38	27 11 1	3.91	3.82	3.76
39	28 12 2	3.65	3.64	3.37
40	29 13 3	3.92	3.89	3.83
41	30 14 4	3.71	3.50	3.34
42	31 15 5	3.57	3.59	3.67
43	32 16 6	3.96	3.96	3.89
44	33 17 7	3.84	3.56	3.60
45	34 18 8	3.96	3.95	3.94
46	35 19 9	3.95	3.94	3.91
47	36 20 10	3.94	3.85	3.87

TABLE 5-1 (Continued): Subtest W2.

Item Numbers		Level		
	(11X12X13)	11	12	13
41	21 9 1	3.97	3.87	3.93
42	22 10 2	3.81	3.76	3.67
43	23 11 3	3.94	3.87	3.72
44	24 12 4	3.91	3.99	4.00
45	25 13 5	3.95	3.97	3.96
46	26 14 6	3.63	3.64	3.75

TABLE 5-1 (Continued): Subtest W3.

	tem Numbers	Level		
	(11X12X13)	11	12	13
67	41 25 1	3.34	3.07	2.78
68	42 26 2	3.00	2.75	2.68
69	43 27 3	3.92	3.77	3.67
70	44 28 4	3.93	3.92	3.97
71	45 29 5	3.34	3.38	3.33
72	46 30 6	3.25	2.99	3.10
73	47 31 7	3.80	3.56	3.57
74	48 32 8	3.79	3.78	3.75
75	49 33 9	3.37	3.13	3.22
76	50 34 10	3.20	3.32	3.24
77	51 35 11	3.63	3.38	3.36
78	52 36 12	3.92	3.76	3.83
79	53 37 13	3.71	3.63	3.57
80	54 38 14	3.49	3.46	3.53
81	55 39 15	3.41	3.06	3.18
82	56 40 16	3.71	3.54	3.44

TABLE 5-2

Three Different Values of Index k* for Each of the Five Alternative Items That Were Administered to All Three Levels, with Their Respective Item Numbers.

Subtest	L1
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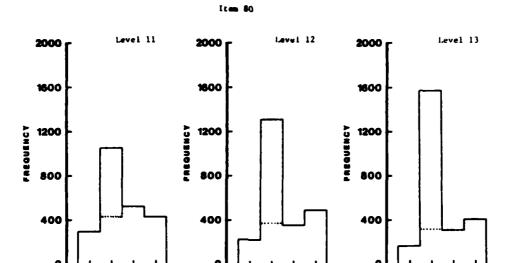
It	em Numbers		Level		
	(11)(12)(1	3) 11	12	13	
58	35 19 1	2.5	3 2.36	2.39	
59	36 20 2	3.54	3.03	2.92	
60	37 21 3	4.1	7 3.99	3.87	
61	38 22 4	3.79	3.62	3.33	
62	39 23 5	2.90	2.83	2.67	
63	40 24 6	4.31	4.14	4.00	
64	41 25 7	3.46	3.28	3.20	
65	42 26 8	3.66	3.36	3.31	
66	43 27 9	4.36	4.32	3.94	

R-88, W1-45, W1-46, W2-44, W2-45 and W3-70, have three estimates of Index k* all of which are greater than, or equal to, 3.9. If we shift this critical value from 3.9 to 3.8, these seven four-alternative test items are joined by eleven more items, i.e., V-63, V-66, R-80, R-90, R-92, L2-58, L3-49, W1-40, W1-43, W1-47 and W2-41. There are no five-alternative test items of Subtest L1 which are comparable to these eighteen four-alternative test items.

We have selected Subtests R and W3, which have the two largest numbers of shared test items, as our examples, and drawn the set of three histograms, which are similar in nature to those in Figures 4-4, 4-5 and 4-6, for the set of three frequency distributions for each item. Figures 5-1 and 5-2 present the resultant nineteen sets of histograms for Subtest R, and the sixteen sets for Subtest W3, respectively.

It is interesting to note that some items show evidence of differential information provided by separate wrong answers.

For example, alternative 4 of R-80 seems to attract students of intermediate reading ability, while alternative 1 of the same item appears to attract students of lower levels of ability. It is clear that many items have one or more effective distractors, and, among others, alternative 2 of R-86 proved to be powerful. Most histograms have some regularities in the way the frequencies change across the three levels, which suggest that the examinees selected their answers intentionally rather than by random guessing.



2

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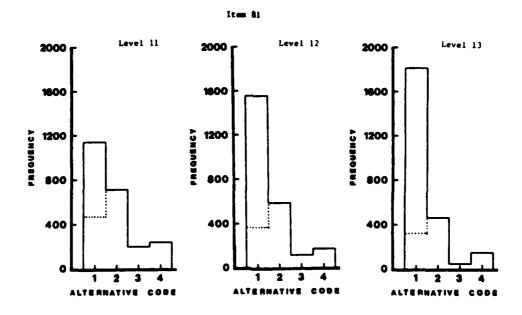


FIGURE 5-1

Comparison of the Three Frequency Distributions of Examinees with Respect to Their Choices of Alternatives for Each of the Twenty Items of Subtest R, Which Were Administered to All Three Levels of Students, with the Estimated Proportion of Examinees Guessing Correctly (Dotted Line).

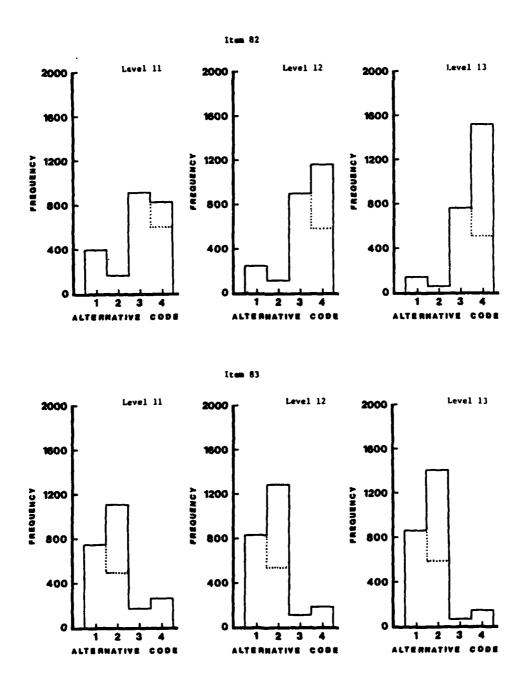


FIGURE 5-1 (Continued): Subtest R.

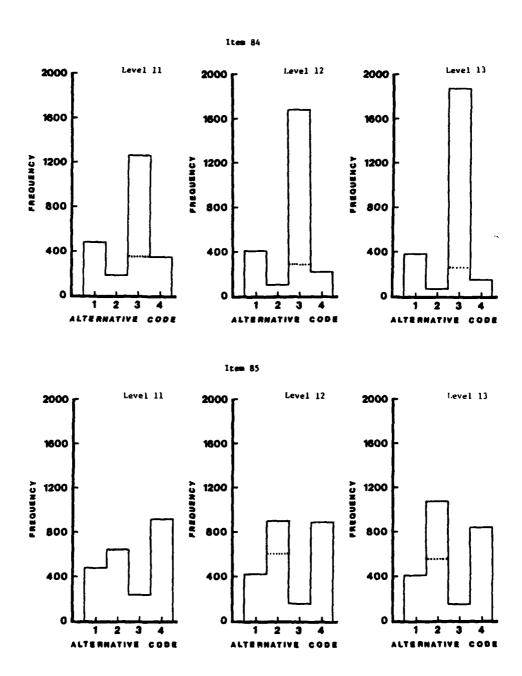


FIGURE 5-1 (Continued): Subtest R.

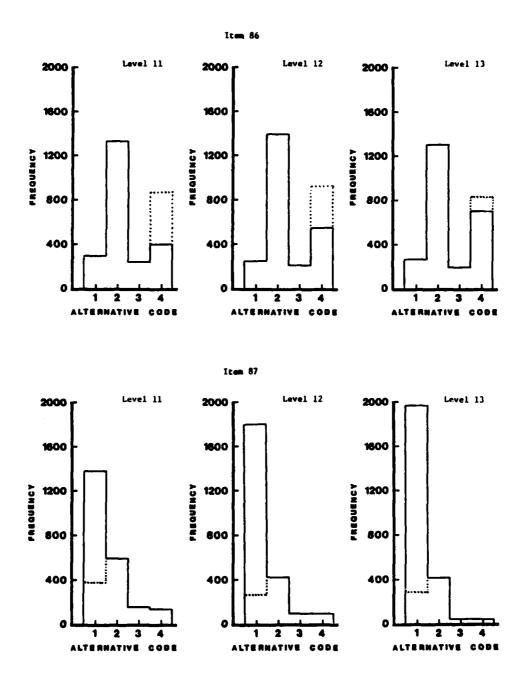


FIGURE 5-1 (Continued): Subtest R.

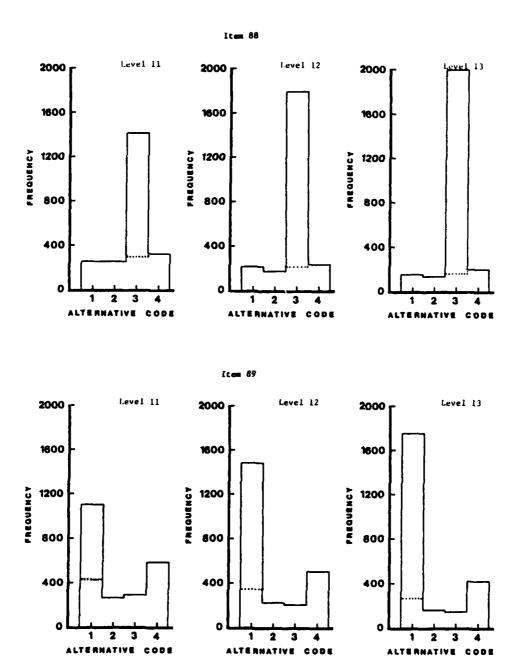


FIGURE 5-1 (Continued): Subtest R.

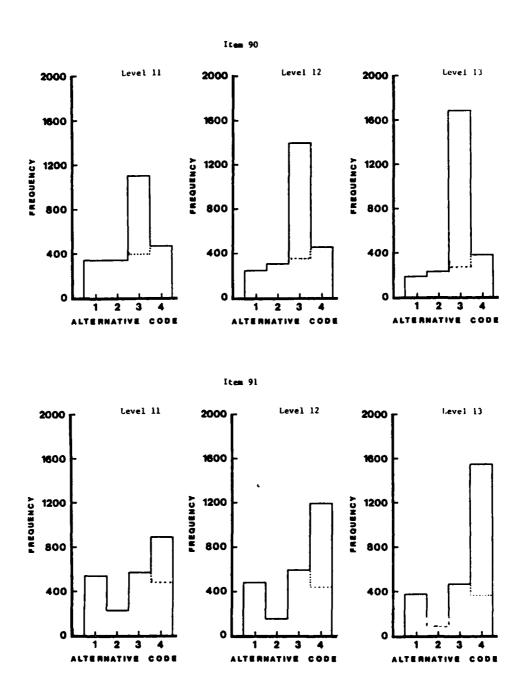


FIGURE 5-1 (Continued): Subtest R.

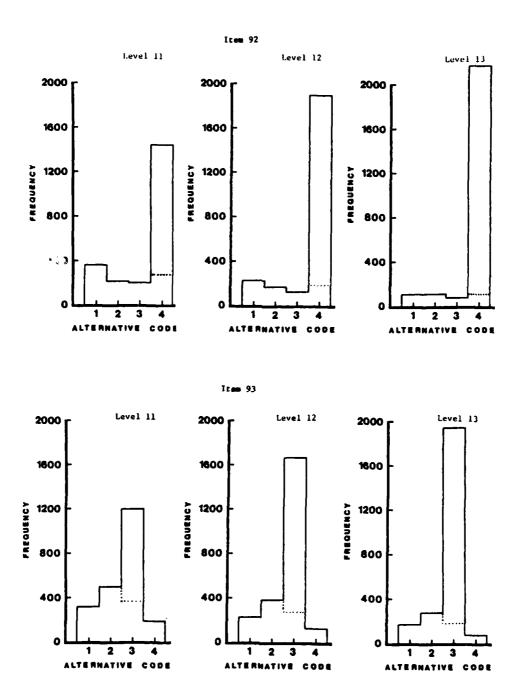
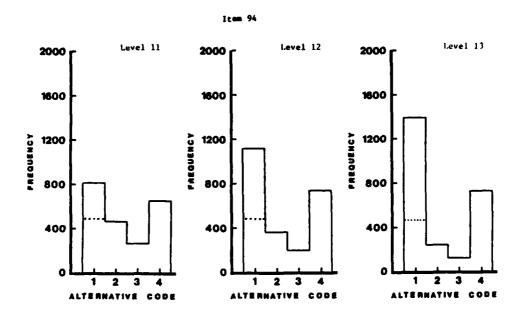


FIGURE 5-1 (Continued): Subtest R.



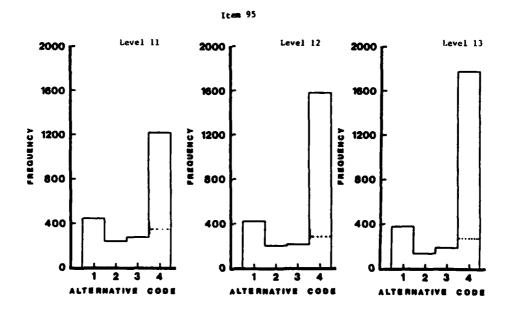


FIGURE 5-1 (Continued): Subtest R.

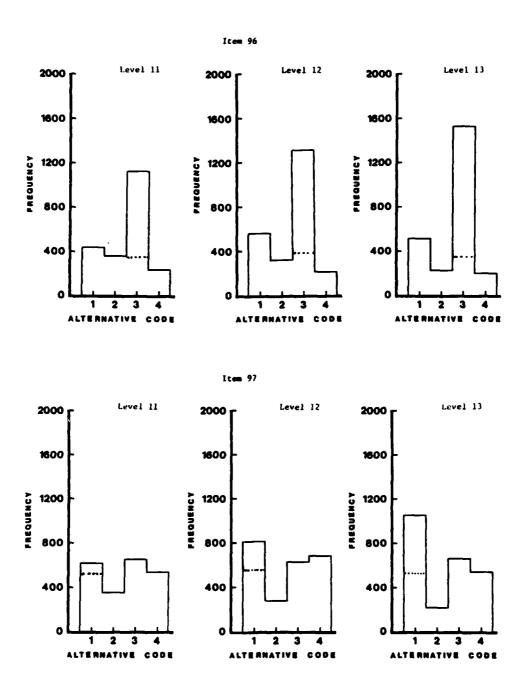


FIGURE 5-1 (Continued): Subtest R.

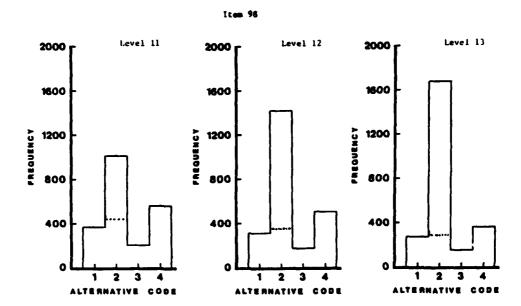
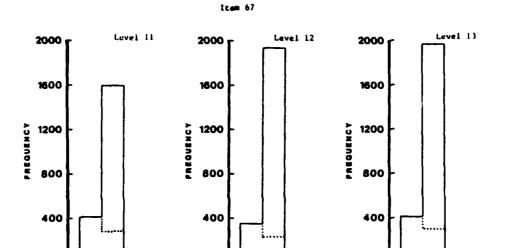


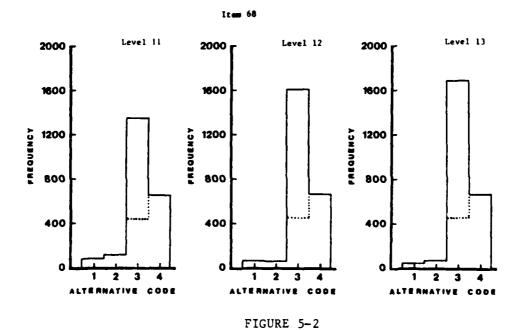
FIGURE 5-1 (Continued): Subtest R.



2

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2



Comparison of the Three Frequency Distributions of Examinees with Respect to Their Choices of Alternatives for Each of the Sixteen Items of Subtest W3, Which Were Administered to All Three Levels of Students, with the Estimated Proportion of Examinees Guessing Correctly (Dotted Line).

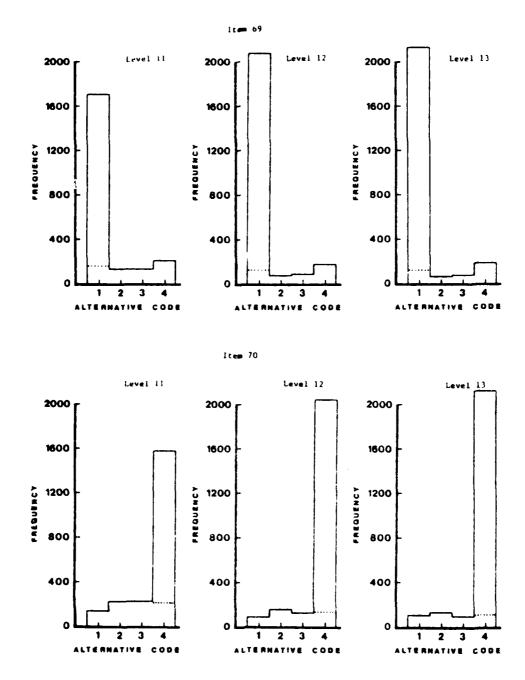
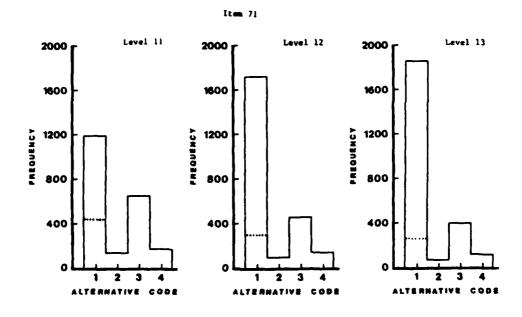


FIGURE 5-2 (Continued): Subtest W3.



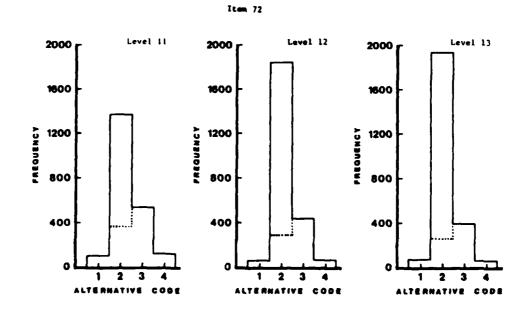


FIGURE 5-2 (Continued): Subtest W3.

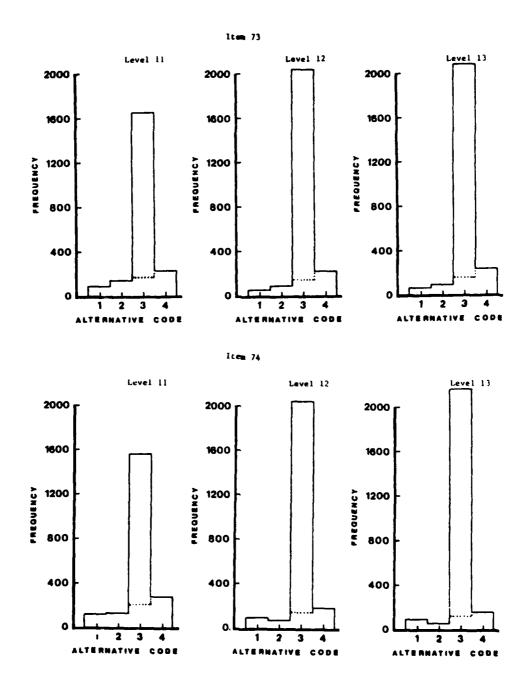
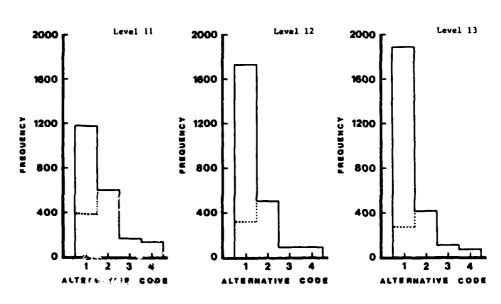
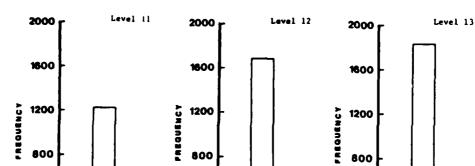


FIGURE 5-2 (Continued): Subtest W3.







400

3

400

Item 76

FIGURE 5-2 (Continued): Subtest W3.

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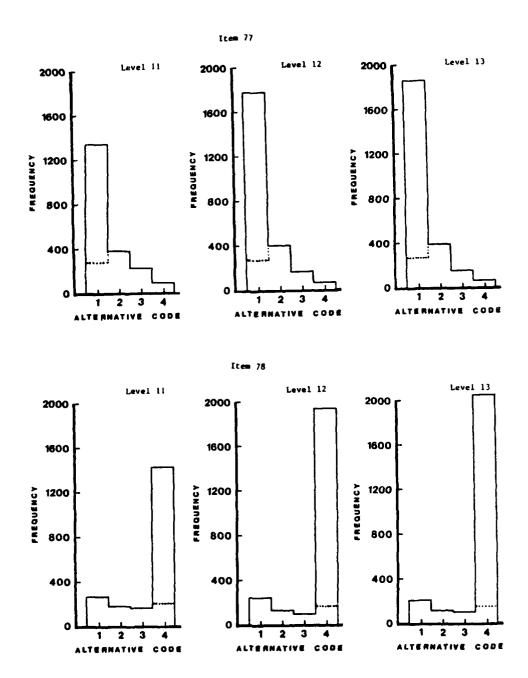


FIGURE 5-2 (Continued): Subtest W3.

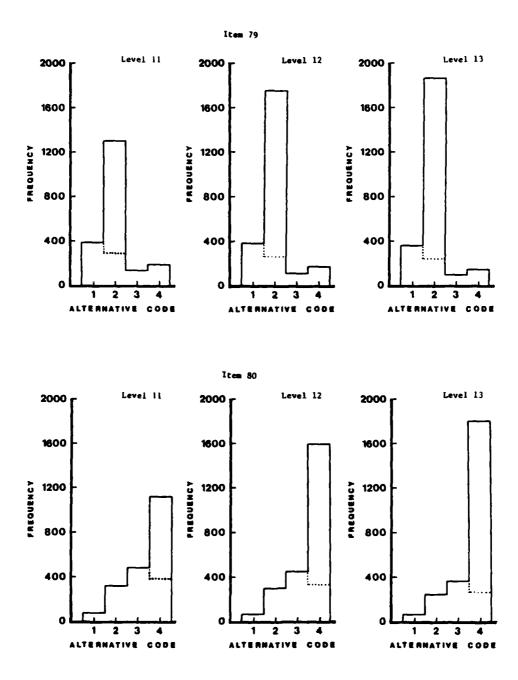


FIGURE 5-2 (Continued): Subtest W3.

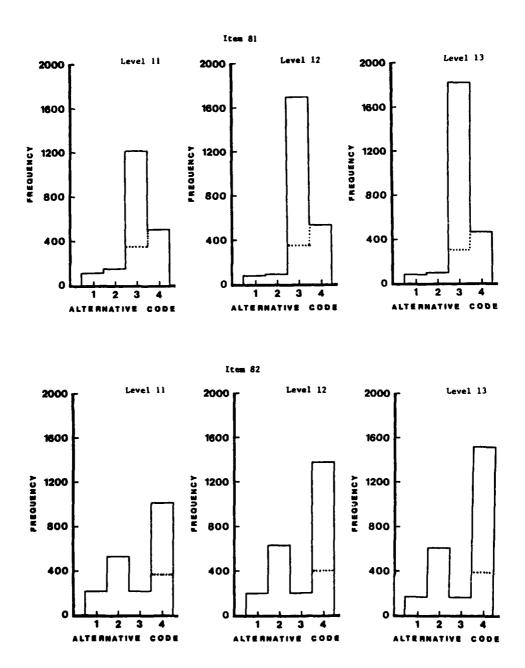


FIGURE 5-2 (Continued): Subtest W3.

VI Discussion and Conclusions

We have seen in the previous chapters that, generally speaking, evidence indicates the direction of Informative Distractor Model, rather than that of Equivalent Distractor Model. This is especially true with the test items for measuring language skills, i.e., those of Subtests L1, L2, L3 and L4. There are certain test items, however, for which the direction of Equivalent Distractor Model may be more suitable, as we have discussed in the preceding two chapters. For these reasons, it will be wise to take the strategy of adopting theory and method which make it possible for us to handle the test items following either general direction.

In so doing, perhaps the most promising way is to adopt the methods and approaches for estimating the operating characteristics without assuming any mathematical form. The selection of items which can be used as a substitute for the Old Test will be the most crucial point in that process. If it succeeds, then we shall be able to conduct the item analysis in the true sense of the word, which leads to the evaluation and suggestions for improvement of each test item.

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APPENDIX I

Bibliography of Selected Research Dealing with the Theoretical Development, or Applications of, the Three-Parameter Normal Ogive, or Logistic, Model

(1) Selection Process

A search was made of some 10 journals, and of the technical reports received by the principal investigator since 1977, for attitles concerning the three-parameter normal ogive, or logistic, model. The selection procedure for these journals was somewhat arbitrary, based on sources cited in the present work, as well as the authors' knowledge of journ is likely to contain such research. The names of those journals and the span of years, volumes, and numbers searched for each journal are given in Table A-1-1.

The articles, reports, and texts found are here organized into applied and theoretical research categories. The applied research categories are paper-and-pencil testing and computerized adaptive testing; the more theoretical works are grouped into areas of general review and development, considerations of the three-parameter approach, Bayesian estimation, and simulation studies. Each has been briefly summarized as it pertains to the three-parameter latent trait model: this was not an easy task.

TABLE A-1-1

Journal Titles and the Respective Span of Years, Volumes, and Numbers Searched for Three-Patameter Model Research Papers

74] 74]	Year, Volume, Number
Journal Hite	Initial Final
American Educational Research Journal	1975,12,1 - 1980,17,1
Applied Psychological Measurement	1977, 1,1 - 1980, 4,2
British Journal of Machematical and Statistical Psychology	1974,27,1 - 1980,33,1
Educational and Psychological Measurement	1965,25,1 - 1980,40,1
Journal of Applied Psychology	1975,60,1 - 1980,65,2
Journal of Educational Measurement	1974,111,1 - 1980,17,1
Journal of Educational Statistics	1976, 1,1 - 1980, 5,1
Journal of Mathematical Psychology	1974,11,1 - 1980,21,2
Psychometrika	1965,30,1 - 1980,45,1
Review of Educational Research	1974,44,1 - 1980,50.1

() Applied Research: Paper-and-Pencil Testing

Lord, F. M. An empirical study of item-test regression. Psychometrika, 1965, 30, 373-376.

On a group of 103,275 examinees who each took a 90 item Verbal and 60 item Mathematical SAT, Lord examines the item-test regressiona. He finds, among other things, that examinees with low observed scores may score significantly below chance levels when taking difficult items. Lord suggests that distractors may be working in these cases.

Lord, F. M. An analysis of the verbal scholastic aptitude test using Sirnbaum's three parameter logistic model. Educational and Psychological Messurement, 1968, 28, 989-1020.

The Item responses of 2,362 examiness who took the Verbal SAT were analyzed using the three parameter model. Efficiency of the model in terms of accuracy of estimation was found to be relatively high, and comparisons of information functions based on subtests were made,

Lord, F. M. Item characteristic curves estimated without knowledge of their mathematical form-A confrontation of Birnbaum's logistic model. <u>Psychometrika</u>, 1970, 35, 43-50.

Five items from the Verbal SAT, which were administered to 103,275 examines, are taken as examples, and the author's method of of estimating the true test score distribution for the given set of observed test mores is applied for both success and failure groups. Based upon these two estimated true score distributions, the estimated item characteristic function is obtained as the taxis of the density for the success group over the total density, and the subsequent transformation of the true test score to the latent ability. When compared to three-parameter logistic curves generated for these items, a close fit between the two was found.

Lord, F. M. Power scores estimated by item characteristic curves.

Educational and Psychological Messurement, 1973, 33, 219-224.

Using information from a re-administered 90-item Verbal SAT aubtest that was originally mistimed for 21 examiness (they were originally tested under far too speeded conditions). Lord estimated ability levels of the examiness.

Lord, F. M. Estimation of latent ability and item parameters when there are omitted responses. <u>Psychometrika</u>, 1974, 39, 247-264.

Using the three-parameter model, Lord analyzed 3 data sets (2926 examinees on a 90 frem verbal aptitude test; 994 examinees on a parallel form of that test; 2946 examinees on an 85 frem arithmetic reasoning test) with a formula-scoring method to maximize the available information.

Jensema, C. J. An application of latent trait mental test theory.

<u>British Journal of Mathematical and Statistical Psychology</u>, 1974,

<u>27, 29-48.</u>

7, 29-48.

Item responses of 4950 examinees to 94 mathematics questions in the Mashington Pre-College Test battery were used to test several hypotheses. Results indicated that (1) Bayesian tailored testing does not lead to substantial reduction in Item number in realistic attuations; (2) parameter estimation might be accomplished graphically; (3) item banks should have vide difficulty ranges and have highly discriminating trans

Lord, F. M. The 'Ability' scale in item characteristic curve theory.

Psychometrika, 1975, 40, 205-217.

Using 6 data sets (2986 V-CRE, 2862 V-SAT, 2926 V-SAT, 2946 H-SAT, 2848 H-SAT, and 1825 Vocabulary WIT at the sixth grade level), Lord found significant correlations over items between difficulty and discrimination parameters. He then proposed a transformation that would eliminate this problem.

Marco, G. L. Item characteristic curve solutions to three intractable testing problems. Journal of Educational Measurement, 1977, 14,

Using some simulation data in conjunction with 1209 CLEP and 1260 APP calculus test examinees, Marco applied the parameter estimation program LOGIST to the tasks of designing a multi-purpose test, evaluating a multi-level test, and equating a test on the basis of pretest statistics. Information functions are applied to the two initial problems.

Reckase, M. D. Unifactor latent trait models applied to multifactor tests: Results and implications. Journal of Educational Statistics, 1979, 4, 207-230.

Uning 1126 examinees who took verbal and quantitative subtests of the Missouri Scholastic Apritude Test, and 1,000 simulated examinees with specific factor-estructured hypothetical tests, Reckase found that (1) one-and three-parameter models estimate different factors when the test has independent factors, but estimate the first principal component when it is relatively large; (2) when independent factors exist, the three-parameter model estimates only one of them, ignoring the others, while the one parameter model seems to estimate the sum of the factors.

Trabin, T. E. and D. J. Weiss. The person response curve: Fit of individuals to item characteristic curve models. <u>Office of Naval Research</u>, Research Report 79-7, 1979.

Using the responses of 151 college students to 216 vocabulary test items, the authors were able to show that three-parameter logistic person tesponse curves were good predictors of test response profiles for thair date in more than 90% of the cases.

(1) Applied Research: Computerized Adaptive festing

out, W. R. and M. D. Reckase. A live tailored testing comparison study of the one- and three- parameter logistic models. <u>Office of Naval Research</u>, Research Report 78-1, 1978.

Vocabulary test items were administered in an adaptive manner by computer to 12% college students. Computeson of relative efficiency curves, test-recest reliabilities, goodness of fit of the models, convergence rates, and eriterion validity indicated the superiority of the three-parameter model over the one-parameter model over the one-parameter model.

3c)if. I. I. and D. J. Weiss. A construct validation of adaptive achievement testing. <u>Office of Naval Research</u>, Research Report 78-4, 1978.

laing two independent groups of 269 and 230 college students, the sauthors administered two stradaptive and two paper and pencial tests of achtevement and vocabulary. Construct validities of the two modes of presentation were approximately equal, with the stradaptive tests scored by maximum likelihood estimation using the three-parameter logistic model requiring 25% to 31% fewer items.

Koch, W. R. and M. D. Reckase. Problems in application of latent trait models to tailored testing. Office of Naval Research. Research Report 79-1. 1979.

Giving a counterbalanced presentation of adaptive and paper-and-pencil achievement tests to 110 college students, the authors found that neither the one- nor the three-parameter logistic model scoring of the adaptive tests yielded satisfactory content validities. Further, the reliabilities of the adaptive tests usere lower than the comparable paper and pencil test presentation of the multidimensional arhievement test.

McKinley, R. L. and M. D. Reckase. A successful application of latent trait theory to tailored achievement testing. Office of Navai Research, Research Report 80-1, 1980.

In a replication of the above study corrected for improper item linking and selection procedures, the authors tested 88 college students with a multidisensional achievement test in an adaptive and a paper-and-pentil counterbalanced test-retest procedure. This time, they found that the one- and three-perameter tests had higher reliabilities over the same material than did a paper and higher rest. The three-perameter model yielded, reasonably enough, higher test information than did either the one parameter model adaptive to the paper and pencil test. Neither adaptive procedure, nowever, yielded satisfactory content validity.

(4) Theorecical Research: General Review or Development

Baker, F. B. Advances in Item analysis. Review of Educational Research, 1977, 47, 151-178.

Baker discusses and reviews research in latent trait estimation, including maximum likelihood approaches, for dichotomous and polychotomous scoring techniques, and describes developments in various models, including the Rasch and Three-Parameter cases.

Hambleton, R. K. and L. I.. Cook. Latent trait models and their use in the analysis of educational test data. <u>Journal of Educational Measurement</u>, 1977, 14, 75-96.

The authors discuss the utility of various latent trait approaches to educational situations, with emphasis on logistic models, information functions, and parameter estimation procedures.

Lord, F. M. Practical applications of item characteristic curve theory. Journal of Educational Measurement, 1977, 14, 117-138.

Lord discusses procedures for designing, or redesigning, a test for latent trait use with the aid of information functions; item blas, two-stage and adaptive testing, and test equating are also discussed and referenced.

Urry, V. W. Tallored testing: A successful application of latent trait theory. <u>Journal of Educational Neasurement</u>, 1977, 14, 181-196.

This paper discusses the necessary conditions for economical and efficient testing of uni- or maliti-dimensional abilities with tailored testing, and describes possible applications to guidance and computer-sided instruction.

Humbleton, R. K., Swaminathan, H., took, L. L., Eikmor, D. R., and J. A. Gifford. Developments in latent trait theory: Models, technical issues, and applications. Review of Educational Research, 1978, 48, 467-510.

In a thorough review of the ffeld, the authors describe the tail range of models employed or developed, various parameter estimation procedures (including maximum likelihood and Bavesian), the use of information tunctions, and various application, i can models. They note in reference to the guessing parameter of the three-parameter model, that stability of the parameter can be found only when the examinee population is quite heteroxenous for the ability.

Warm, T. A. A primer of item reponse theory. <u>Department of Iranaportation</u>, Tachnical Report 941078, 1978.

This primer gives a general, expository presentation of latent trait theory and specific techniques of application, suggesting use of available computer programs where available and giving computation procedures classwhere.

Lurd, F. N. Applications of item response theory to practical testing problems. Hillsdale, New Jersey: Lawrence Erlbaum Assoc., 1980.

In this text, Lord presents a wide-ranging coverage of most copics involved in the adaptation of latent trait theory to real testing situations. He integrates and updates his past research, as well as that of most others in the field, and includes rationales and justifications for various applications and techniques.

(5) Theoretical Research: Considerations of the Three-Parameter Approach

Samejima, F. A general model for free-response data. Paychometrika Monograph, 1972, No. 18. Samejima integrates the dichotomous response level into her general model for free-response data, and shows that, since the three-parameter model has a lower asymptote greater than zero for the correct-answer category, the basic function cannot be strictly decreasing in ability, and, therefore, the unique naximum condition cannot be satisfied. This implies that there may not exist unique maximum likelihood eatimates for some response patterns.

Samejimm, F. A comment on Rimbaum's three-parameter logistic model in the latent trait theory. <u>Psychometrika</u>, 1973, 38, 221-233.

Based upon the fact that Birnbaum's three-parameter logistic model does not assure a unique maximum for the likelihood function for every response pattern, the critical value, \$\int_{0}\$, of shility \$\theta\$ for each item \$\pi\$ is introduced. It is suggested that item \$\pi\$ should not be used for estimating the examines's ability below this point.

van der Linden, W. Porgetting, guessing, and mastery: The Macready and Dayton models revisited and compared with a latent trait approach. Journal of Educational Statistics, 1978, 3, 305-317. Comparisons are drawn and distinctions made by the author in a comparison of Cuttamn-like mastery madels with the three-parameter logistic model. The latent trait model is extended with the conception of mastery as a region on a latent variable.

(6) Theoretical Research: Bayesian Estimation

consens, C. 1. The validity of Bayesian tailored cesting. Educational an Psychological Mesurement, 1974, 54, 757-766.

Using sets of 100 Monte-Carlo simulaten examinets on iour different item pools, Jenseas found that the standard error of the estimate was a more accurate criterion for terminating item presentation than was number of items presented; the only exception to this generalization occurs when the item discrimination parameters roughly equal the guessing

McBride, J. R. Some properties of a Bayesian adaptive ability testing strategy. Applied Psychological Measurement, 1977, 1, 121-140.

In a series of 4 studies, with finite or infinite item pools and with normally distributed population or distribution-free context criteria. Weside examines the characteristics of Owen's Bayesian procedures for adaptive testing on simulated data. We finds a high correlation between the estimate of ability and the true ability, but a lato finds that the estimate at a positively related to number of presented items, is bissed, and waries is accuracy with respect to ability.

Theoretical Research: Simulation Studies

3

Nambleton, R. K. and R. E. Iraub. Information curves and efficiency of three logistic test models. <u>British Journal of Mathematical and Statistical Psychology</u>, 1971, 24, 273-281.

Using guessing-free simulation data, the authors show that the relative efficiency of the one-, tow., and three-parameter models, as defined as the ratio of information functions between two of them, is quite close if the range of item discrimination parameters is restricted, regardless of ability level. With guessing introduced, the three-parameter model is the much more efficient at lower ability levels.

Lord, F. M. Robbins-Honro procedures for tailored resting. Educational and Psychological Measurement, 1971, 31, 3-31.

Lord demonstrates the use of a ahrinking-step-size procedure for salection of subsequent Iseas in tailored testing, using a theorestical item pool with known item parameters for the three-parameter logistic model. We shows advantages of this procedure over fixed-step-width procedures, but indicates that item pool size increases exorbitantly in this approach if more than six or seven items are t. be administered per examines.

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Lord, F. M. A theoretical study of two-stage testing. Psychometrika, 1971, 16, 227-242.

Using 100 different simulated two-stage testing designs, Lord finds some of them superior to adaptive testing, if guessing is not possible in the first stage of testing. Where guessing surfacely occurs 20 percent of the time or more on the first stage, no two-stage procedure is as effective as tailored testing that can adapt up and down the ability spectrum.

Urry, V. W. Approximations to Item parameters of mental test models and their uses. Educational and Psychological Measurement, 1974, 34, 253-269.

the lies parameters in the initial approximation of the lies parameters in the three-parameter models for use in the screening for adaptive testing and as first values for maximum likelihood estimation. The guessing parameter of the item-test regression curve. This parameter of the item-test regression curve. This parameter is then used to convert point-biserial item-test correlations to discrimination estimates and proportion correct to difficulty estimates.

iensema, C. A staple technique for estimatine lutent truit mentul test parameters. <u>Educational and Psychological Measurement</u>, 1976, 36, 705-715,

Fullowing Urry's lead, lensema describes straight-forward techniques to estimate three-parameter model parameters. In comparison with maximum likelihood estimates, based on 2,26 simplicated items, the simplie estimates correlated 0.86 for the discrimination parameter and 0.97 for the difficulty parameter. Accuracy of the technique increases as sample size and increase, and decreases as item discrimination increase.

Lord, F. H. Optimal number of choices per item -- A comparison of four approaches. fournal of Educational Measurement, 1977, 14, 33-38.

Under various simulated conditions, 90 items of the Verbal section of the SAT are manipulated such that their guessing beneters are either left as is, or changed to conform to 5-4, 1-, or 2-choice items. Lord finds ability-specific results for these conditions. Decreasing the effective number of choices per item, while lengthening the test in proportion, who guess little if at all, but decreases it for low-ability examinees, who might be expected to guess more frequently.

Schmidt, F. L. The Urry method c. approximating the Item parameters of latent trait theory. Educational and Psychological Measurement, 1977, 37, oll-620.

Schmidt shows that Urry': (1974) method of parameter estimation underestimates the discrimination parameter, and overestimates the difficulty parameter. He also demonstrates a simple procedute to correct these biases by transforming the point-biserial correlation.

Levine, M. V. and D. B. Rubin. Measuring the appropriatences of multiple-choice test scores. <u>Journal of Educational Statisties</u>, 1979, 4, 269-290.

Using several indices of internal consistency and three-parameter model assumptions, Levine and Rubin explore simulated examinee item response patcerns to SAT questions. They find that sufficient information does exist on the answer sheet alone to judge the appropriateness of the test for a given examinee or group of examinees.

Ningsbury, C. C. and D. J. Weiss. Relationships among achievement level estimates from three Item characteristic curve scoring methods. Office of Naval Research, Research Report 19-1, 1979.

The authors compare maximum likelihood normal, maximum likelihood logistic, and Bayesian scoring methods using one-, two-, and thee-parameter models for various data sets. The results indicate that for the three-parameter model, Bayesian estimation produces results divergent from the maximum likelihood techniques, but does not suffer from problems of non-convergence; convergence problems are maximal for maximum likelihood procedures with convertional tests of inappropriate difficulty.

Rev. M. J. Estimating item characteristic curves. Applied Psychological Measurement, 1979, 1, 171-185.

Ree uses 2,000 simulated subjects' responses to BO-item tests to judge the effectiveness of 4 different computer estimation procedures, where the distribution of shility is rectangular, random, or normal. OGUNA was found to work best with the latter data set; LOGIST on the random distribution works must accurately, but costs substantially more in computer time.

APPENDIX II

Directions of Iowa Tests of Basic Skills for Each of the Eleven Subtests

PERCYTONS FOR LACH OF THE ELEVEN STRIFS.,

The same of the sa

(I) Subtest V

In each exercise, you are to decide which one of the four answers has must nearly the same meaning as the word in heavy type above them. Then, on the answer same as the exercise you are working on. You are to fill in the inswer space on the answer the same as the inswer space on the answer sheet that has the same maker as the inswer sheet that has the same to pricked. The sample exercise in the box at the right has already been marked currently on it answer sheet.

Sample Exercise Close the door

2) hold 9) behind 4) open

Answer 11

Subtest A

This test consists or several reading selections. After each selection there are some exercises. Reading selection quickly and then answer the usertises, four answers are given for each evertise, but only one of these answers is right. You are to choose the one answer that you think is better than the orders space. Then, on the answer sheef, find the row of answer space. Spaces numbered the same as the exercise. Fill in the answer space for the best answer. The sample exercise shell whom shows you how to mark your answers on the answer.

Sample Exercise: Every Sunday after linner Eup gets a ball Rame on D., The next ching

a ball game on D. The next the vertice of the bound he is snoring.

What does Pop do on Sunday afternoon?

1) Morks in the yard

2) Yoes to church,

3) Takes a nap

(answer 3)

1

(1) Subtest [1]

The exercises in this spelling test are libe the samples shown at the right. Many of the exercises contain a mistake in spelling. Some do not have any anstakes at all. You are to look for mistakes in spelling. When you find a mistake, fill in the enser space on the answer sheet that has the same mistake in an exercise, fill in the mistake in an exercise, fill in the first has the same mistake in an exercise, fill in the fifth answer bear. The sample exercises at the right show our what the contract of the sample exercises at the right show our

diplic fact. 15cs

(4) Subtest [2]

This is a test on capitalization. It will show whether you know which words in a sentence should be apitalized. The exercises in the test are like the samples shown below. Many of the exercises on the exercises contain mistakes in capitalization. Some who not have any mistakes in capitalization. Some who not have any mistake at all. You are to look for a mistake, if ill in the masure space on the answer the mistake. If there is no mistake in an everyable that has the same number as the line containing the mistake. If there is no mistake in an everyable intil in the fourth answer space, the sample contains below show what you are to do.

Sample Exercises:

S1. 1) Tow and jerry
2) picked up all the
3) trash from the picht.
4) thus from the picht.

- 1) Sally said that 55
- 2) everyone should have 3) been more Careful. 4) (No mistakes)
- (answer: 3)
- 1) Let's all help
 2) to keep our streets
 3) and sidewalks clean.
 4) (No mistakes) 3.3
- [(5 :Jansue]

(5) Subtest L!

you can use periods, commas, question marks, apostrophes, etc. The exercises in the test are like the samples in punctuation. Some do not have any mistakes at all. You are to look for mistakes in the test exercises. Then you find a mistake, fill in the answer space on the answer sheet that has the same number as the line containing the mistake. If there is no mistake It will show how well shown below. Yany of the exercises contain mistakes he sample exercises below show what you are to do. in an exercise, fill in the fourth answer space. This is a test on punctuation.

ample fxercises:

- 1) Our family tries 2) to practice . S
 - rules of safety rules of safet
 (No mistakes)
- [answer: 🖰]

 - (answer: 3) 1) We all fasten
 2) our seat belts
 3) before, we leave.
 4) No mistakes)
- ie do our best
 to male our home
 a sate place to live.
 io mistakes) {anstakes} . آد
- [answer: 4)]

(f.) Jublest Li

This is a test on the use of words. It will show whether we know how to use acids according to the standards of cerrorith critical English. The exercises

that has the same number as the line containing the mistake. If there is no mistake in an exercise, fill in the fourth answer space. The sample exercises below show what you are to do. of the exercises contain mistakes in the use of words. Some do not have any mistakes at all. You are to look for mistakes in the test exercises. Then you find a mistake, fill the answer space on the answer sheet in the test are like the samples shown below.

Sample Exercises:

- 1) He showed us the way. . S1.
 - Are you afraid to try?
 Me and him took turns.
 (No mistakes)
- [answer: 3)]
 - Time vent first.
 The bird flew away.
 Pat found a dollar.
 (No mistakes) \$2.
- (Answer: 4)

Subtest W1

2

these answers is right. You are to choose the one answer that you think is better than the others. Then, on the answer sheet, find the row of answer spaces numbered the same as the exercise. [11] in the answer space for the best answer. This is a test of your ability to read maps. It contains several maps, with some exercises about each one. Four answers are given for each exercise, but only one of

(8) Subtest 1.2

This is a test of your ability to read graphs and tables. After each graph or table there are several exercises. For each exercise, decide which answer is correct. Then mark the proper answer space on the answer steet. Mark only one answer space for each exercise.

Subtest 177

6

This is a test of study skills such as looking up words, alphaketizing, using an index, and locating information. Read the directions for each part carefully, and then mark your answers to the exercises on your answer sheet.

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Visit of the alternation of the control of the following the second of the control of the contro
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was the contract of the text are like the text are an exercise and a war to the first the like the like and the text are th

JB, 1 Ker , se

James Parker of the How many benefit and the crees, flow many benefits and sixters, does she happed. Permittan District Less III all and the second of the seco

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APPENDIX III

Relative Frequencies of Examinees Who Left the Items
Unanswered in Each Subtest, Arranged in the Order
of Presentation of the Items

TABLE A-3-1

Relative Frequencies of Examinees Who Left the Items Unanswered in Each Subtest, Arranged in the Order of Presentation of the Items: Level 11.

Subtest V	0.000	8000.0	0.0013	0.0013	0.0025	0.0008	0.0017	0.0021
	1.0017	0.0017	0.0021	30.00	0.0038	0.0317	0.3038	0.0051
	0.0021	0.0030	6.0008	00 00 0	0.0051	0.0030	0.0017	0.0025
	0.3034	3.0355	C.0072	J. JU55	0.0118	0.0113	0.0165	0.0140
	0.0228	C.C271	0.0343	0.0381	90,0.0	0.0444	9650.0	0.0626
	3.0660	0.0740	C.0863					
Subtest R	3.0333	0.0304	C. CC25	0, 3034	0.0013	0.0013	0.0042	0.0008
	7.30.4	J000.J	0000	0.3004	0.000	u. 000 8	3.3004	00000
	1.3346	0000.0	3.0017	0.0013	3.3038	0.0000	0.0013	9000.0
	5.0004	0.000.0	C.0C17	0.0013	3.0000	0.3025	000000	0.0008
	3.3321	6.0313	0000.0	3.3004	0.300B	0.0008	0.0034	0.0017
	0.0021	0.003C	0800 -0	0.0034	0.0034	0.0051	0.0068	0.0106
	3.3308	3.0005	6830.3	0.110	0.0118	0.0135	0.0131	0.0266
	0.0288	0080.0	0.0326	0,0326	0.0364	0.0402	0.0419	0.0453
	3.2480	0.0520	C. C584	0.0047	0.0698	0.0740	0.080s	0.0850
	3.3935	1686.0						
Subtoot 11	ć		4		,	6	0	
	0000	2002.0	5100.0	0.0004	0.000	6.000	0000.0	*000°0
	3.000	0.0013	C. OCCB	0.0008	0.0013	0.0013	0.0013	8000.0
	4.00C • 0	7 (100° 0	£222.7	o. 30 ce	0.0034	0.0021	2.0017	C. 3C.3
	1.3338	3.00.5	0.000.0	3.311)	2.10.0	5.0144	0.0199	0.0258
	2,0305	0.0300	C.0402	0.0486	0.0592	0.0706	0.0825	0.0931
	3.1303	0.1129	0.1273					
Subtest L2	usio.	8,000	0.0013	0. 30 CH	0.0008	0.0008	000000	3.0004
	6100.0	8000°n	0.0017	0.3317	0.0004	0.0008	9000°0	3.0008
	5.3317	3.0013	t200.0	000000	0.0000	0.0030	0.003A	0.0042
	3.3347	3.3072	0.0076	3.0093	0.0097	0.0118	0.3152	0.0178
	3.0237	0.0250	0.0266	0.0368	0.0431	0.0499	0.0571	0.0635
Subtest L3	0.0000	0.000	C.0030	0.0004	0.0043	0.0004	0.0004	0.0004
	3.3304	6.0008	C.0C25	0. 00 30	0.0CJB	0.0004	J.000E	0.0017
	3.3338	3,3038	0.0003	3,3003	0.0038	0.000.0	0.0017	9000
	0.0013	0.0034	CE33 °3	0.0051	0.0051	0.0055	1400.0	8400.0
	3.0085	0.0102	0.0085	0.0089	0.0127	0.0135	0.0144	2.0157
								1

TABLE A-3-1 (Continued): Level 11.

L4 W1	0.0000 0.0000 0.0000 0.0000 0.0005 0.0005 0.0005 0.0005	00000000000000000000000000000000000000	0.0021 0.0021 0.0013 0.0076 0.0008 0.0000 0.0013 0.0351	0.00017 0.0006 0.0076 0.0076 0.0076 0.0076 0.0076 0.0076 0.0076	0.0000 0.0000 0.0000 0.0076 0.0025 0.0025	0.00117 0.0008 0.0004 0.0004 0.0004 0.0034	0.0013 0.0017 0.0097 0.0013 0.0089 0.0089	0.0013 0.0021 0.0017 0.00114 0.00055 0.00056 0.1404
	1.0004 7.0004 0.1047	0,000 0,1000 0,0051 0,0563	C.0004 U.0008 C.0068	0.0004 0.0017 0.0089	0.0004 3.2004 0.0182	0.0004 0.0013 0.0266	0.0008 0.0025 0.0338	0.0013 0.0030 0.0398
	0.0004 0.0017 0.0025 0.0025 0.0066	6.0017 5.0021 6.0008 5.0021 5.0157 6.0719	C.C033 0.0017 C.0030 C.0025 0.0220 0.0774	0. 0025 1. 0021 0. 0034 0. 0025 0. 0246 0. 0846	0.0021 0.0055 0.0013 0.0025 0.0355 0.0897	0.0034 0.0013 0.0013 0.0410 0.0977	0.0030 0.0330 0.0008 0.0089 0.1062 0.1062	0.0008 0.0030 0.0021 0.0102 0.0541 0.1096
	3.0004 3.0110 9.0325 3.0055 0.055	1200.0 1000.0 1000.0 1000.0 1000.0	C. 0CU4 C. 0908 0.0C25 0.0G5)	0.0008 0.0017 0.0008 0.0033 0.0033	0.0055 0.0017 0.0030 0.0030 0.0385	0.0008 0.0017 0.008 0.0157 0.0423	0.0055 0.0030 0.0042 0.0228	0.0013 0.0013 0.0089 0.0152
	0.0000 0.1012 1.0138 1.0575	0.0008	C.0008 C.3008 0.CC65 C.1C03	0.0025 0.0051 0.0152 0.1210	0.0038 0.0008 0.0250 0.1426	0.0055 0.0008 0.0008	0.0004 0.0038 0.0355	0.0013 0.0030 0.0499

TABLE A-3-2

Relative Frequencies of Examinees Who Left the Items Unanswered in Each Subtest, Arranged in the Order of Presentation of the Items: Level 12.

Subtest	>	40.00	000000	0.0341	3.3304	900000	7 100.0	0.3004	9.0008
		0.0025	6.00.04	6000.0	0.0012	0.0021	0.000	0.0029	0.0021
		0.0317	800000	0.0017	3.30.98	J.0C12	7.00.0	0000.0	0.0017
		9.0204	0.0004	0.005)	0.3041	0.0054	0.0053	0.000.0	0.0054
		0.0108	0.0099	C.0C87	0.0112	0.0170	0.0174	0.0240	0.0278
		0.0323	3. (356	0.0431	0.0464	0.0481	0.0510		
Subtest	æ	0.000	6.0004	8000.0	7100.0	0.0012	2.000.0	0.3008	9.0029
		3, 3038	0.3033	C.0074	3,0625	0.0004	0.000	0.0000	4000.0
		3,3021	033	8000	C. U120	0.0000	\$ 00° 0	0.3021	0.0017
		3, 3029	0.0029	0.0012	7100.0	0.0021	7 100.0	4000°0	0.0004
		3.0012	0.0021	0.0017	0.0037	0.0621	0000.0	0.0008	0.0033
		3.3004	6.0308	C.0C29	0.0017	0.3029	0.0033	0.3054	9,00.0
		0.0054	0.0116	C. 0128	0.3157	0.0136	0.0170	0.0199	0.0236
		0.0240	0.0286	C.0269	0.0274	J.0369	0.0414	0.0435	0.0506
		0.0526	C.C555	C. C583	0.0613	0.0663	0.0080	0.070.0	0.0738
		3.0796	9619.0	0.0821	0.0633				
Subtest	-		90000	6	7600 0	6100	ò		
20000		0000		0.0012	1000.0	7100.0	*000.0	2000	0000
		5.00.04	0.0004	t 0000 t	0.0004	0.0004	0.000	8000.0	0.0021
		3.0325	6.0004	0.0000	0.00C4	0.0012	0.0012	0.0012	0.0012
		5.0029		C.C58	J. 00.66	0.0075	\$010.0	0.0128	0.0191
		3.3230	C. 0282	0.0323	0.0344	0.0493	0.0576	0.0688	0.0796
		1.0932	0.1015	C. 1111	U. 12 02	0.1281	0.1359		
Subtest	L_2	ecro*:	0.0004	0.009	0.0004	0.0034	0.0325	3666.0	0.0008
		0.3000	0.0012	0.0012	3.0012	J.0000	0.000	0.0008	0.0012
		2.0017	6.0004	4000°0	0.0000	0000.0	0.0008	0.0012	0.0004
		0.00.0	0.0008	0.0012	0.0033	0.3062	0.0066	0.0087	0.0104
		0.0133	0.0162	0.0224	0.0236	0.0269	0.0323	0.0361	0.0385
		0.0423	6.6435						
Subtest	L3	60.66.0	0.000	C.0004	0.0000	0.0004	0.0017	0.0004	0.0000
		3.000	2,000.2	0.0021	0.0025	0.0004	0000.0	0.0008	0.0012
		3.0012	0.0017	C-0C04	0.3004	2100.0	2.0017	7100.0	3.3021
		0.3008	C.0012	0.0017	0.0012	0.0012	0.0021	0.0021	0.0033
		1.00.1	0.0041	0.0062	0.0058	0.0083	0.0099	0.0120	0.0128
		0.0120	0.0120						

TABLE A-3-2 (Continued): Level 12.

Subtest L4	3.300	6000.0	C. CC12	0, 0000	0.0006	0.000	900000	C. 3029
	0-0004	C. CCO 4	0.0017	0.0041	0.000	0.0008	0.0017	0.0012
	0.0004	0.000	0.0012	0.0012	0.000	0.0004	0.0012	0.0025
	2.0017	0.0037	0.0033	0.0037	0.0017	0.0025	0.0037	0.0050
Subtest Wl	3.3041	0.0025	+000-0	0.0341	0.0012	0.0321	0.0008	40000
	3.3312	40.0.0	C.0C17	3. 30 12	0.0034	0.0003	0.0025	0.0037
	1.3000	C100.0	0.0021	0.0012	0.0017	0.3095	1600.0	0.0124
	0.0199	0.0253	C.0265	ù.0327	0.0365	0.0452	0.0506	0.0739
	1960.0	0.1105	C.1372	0.1625	1181.0	0.2063	0.2366	2,2482
Subtest W2	4000	0.0021	0.0008	0.0004	0.0334	0.0000	0.0012	0.0000
	3.0008	0000.0	c. coc8	0.0012	0.000	0.0004	0.0017	0.0004
	2.0017	9700.0	9,00.0	0.3087	0.0153	0.0207	0.0249	0.0298
	0.0493	0.0634	1110.0	0, 3866				
Subtest 113		,		3000	6	0		
	71000		100.0	6700.0	2000	900000	0.0025	c. 0008
	0000	2022	c. 0004	0.0017	0.0017	0.0012	*000*0	0.0029
	3,0008	0.0012	0000	3.0008	0.0033	0.0021	0.0021	C. 0008
	3.3004	8000.0	0000.0	3.000B	0.0012	0.0012	0.0017	0.0008
	5.3017	0.0025	C. CC21	0.0029	0.0029	0.0033	0.0050	9,0000
	1.3050	0.0010	0.0099	0.0133	0.0145	2.0182	3.017e	0.0224
	92 20 ، د	0.0315	0140.0	0.0410	0.0472	0.0522	0.0539	0.0597
	3.0655	0.0696	0.0750					
Subtest M1	0.000	800000	0.0025	0.0004	0.0017	2100.0	40000	6000
	3, 3325	c. co2 s	0.0025	2.0012	0.0021	0.0012	0.000	0.0008
	0.0008	0000.0	0.0050	0.0008	0.0012	0.0012	0.000	0.0033
	3.0012	0.0033	C.0037	0.0017	0.0075	0.0099	0.0066	0.0091
	0.0145	0.0133	C.0141	5, 01 78	0.0195	0.0220	0.0253	0.0286
	0.0381	0.0448	0.0448	C. 0547	Ú.0555			
Subtest M2	3.0000	0.000.0	00000	0.0037	0.0058	0.0008	0.0029	0.0025
	J.0312	0.0037	C.0038	900000	0.033	0.033	0.0041	00000
	12.00-0	1,00.0	1,20,2	0.3054	0.0104	0.0104	0.0149	0.0220
	0.0274	0.0361	0.0464	0.0535	0.0100	0.0783	0.0974	
	_							

TABLE A-3-3

Relative Frequencies of Examinees Who Left the Items Unanswered in Each Subtest, Arranged in the Order of Presentation of the Items: Level 13.

Subtest V	1336	4100 0	4100	7100	000	0000		
	0.0.16	9(00)	4100	200		0000	00000	*000*6
	3.0316			0.000		000	0.00	9700.0
	3.0008	2000.0		97000	9600	0000	0.0020	2100.0
	1.0.140	1000		0.00	*****	* 200.0	0.0032	0.0032
		***************	66000	7400 · 0	0.000	2000	0.0101	0.0138
	110.0	0.6122	Z*I3*3	0.0158	0.0194	0.0207	0.0203	0.0198
Subtest R	3.3308	0,000	c.000a	0.0008	0.0012	0.0016	0.0024	0.0020
	2.0016	0.0012	6,0003	0.0000	0.0020	0.000	00000	0.0012
	3.00.0	0.0316	0.0032	6.0117	0.3034	0.000	9000.0	0.0012
	3.00.8	C.CC24	C.0012	J. 30 CB	0.0028	0.0012	00000	0.0016
	0.0020	0.0041	0.0024	0.0004	0.0023	0.0020	0.3024	8000.0
	7. 03 CB	C. C034	C. OCC4	0.0008	0.0020	0.0023	0.0016	0.0028
	3.0049	6.0041	C.0026	0.0041	0.0053	0.3061	0.0053	0.0061
	2.0065	c.0105	0.0142	0.0146	0.0122	0.0194	0.0215	0.0235
	0.0239	C.2207	C.0288	0.0304	0.0324	0.3401	0.0346	0.0381
	7.141.1	0.3454	0.0458	0.0470	0.0474	0.3494		• •
Cultipost 11	,	1						
TT 1Salone	0.000	0.0016	2.0012	0.0316	0.0012	0.0008	0.0004	9000-0
	1.330	0.3336	+c00.0	900000	0.0008	0.000 R	0.0020	0.0012
	3.3338	4,000.0	C. OCC8	C. 0C08	0.0624	0.0008	0.0020	0.0004
	5.0000	J.0J36	0.0045	0.0057	0.0089	0.0397	0.0113	0.0130
	0.3166	C. C186	0.0223	0.0243	0.0296	0.0360	0.0413	0.0458
	3.0498	0.0531	0.0628	0.0705	0.0814	0.0871	0.0895	0.0944
Subtest L2	1,00.04	400000000000000000000000000000000000000	K. 00.0	0.0036	6100	2000	6	
	0.0038	000000	0.00.0	0-0027	7100.0		******	4000
	0.0016	4000-11	0000	070000	7000	2000	71000	0.0020
	0.00.0	4700.0	0.0036	0.00.24	7600	100.0	10000	8000.0
	1.0093	0.1176	n. 0177	77 W 10	7714	1 to 20 to 2	60000	2000
_	3.0227	0.0235	0.0292		70.0		0.020.0	0.0633
Cubrost 13								
Sancest E3	7.0004	2000.0	6000	2. 00 CB	9100.0	0.0016	0.000.0	0.0004
	9100.0	0000°0	0.0016	0.0032	0.0004	0.000 B	0.0008	0.0016
	J. 2008	4000.0	C. 0COB	0.0008	0.0016	0.0004	0.0004	0.0008
	3.0304	0.0004	0.0016	0.0016	0.0338	0.0016	0.0012	0.0008
	0.0016	0.0024	C. 0C41	0.0045	0.0036	0.0061	0.0081	0.0077
	1.0385	5800.0	c.c105					
T								

0.0004 0.0004 0.0008 0.0028 0.0227 0.0535

0.0012 0.0004 0.0113 0.0161

0.0008 0.0053 0.0077 0.0680

0.0057 0.0028 0.0065 0.0065

0.0012 0.0008 0.0101 0.0531

0.0004 0.3004 0.0065 0.0425

C.0012 U.0038 O.0C49 C.0316

0.0004 0.0024 0.0045 0.0045

0.0000 0.0012 0.0032

M2

Subtest

0.0012 0.0012 0.0024 0.0150 0.0150

0.3000 0.3028 0.0024 0.3024 0.0122 0.0004 0.0012 0.0008 0.0036 0.0004 0.0032 0.0036 0.0425 0.0024 0.0016 0.0174 0.0024 0.0024 0.0000 0.3045 0.3049 0.0000 0.0000 0.0000 0.0000 0.0012 0.0038 0.0041 0.3271 0.0004 0.0012 0.0000 0.0024 0.0032 0.0041 0.0166 0.0024 0.0012 0.0012 0.0016 0.0093 0.0016 0.0012 0.0012 0.0267 0.1000 0.0000 0.0000 0.0000 0.0000 0.0012 0.3624 0.3020 0.0612 0.3320 0.0028 0.0166 0.0012 0.0034 0.0028 0.004 0.0089 0.0008 0.0049 0.0016 0.0020 0.00249 0.0064 0.0166 0.000% 0.0012 0.0049 0.0441 0.0012 0.0016 0.0016 0.0016 0.0012 0.0142 0.0016 0.002d 0.0020 0.0004 0.0085 C.0C23 C.0004 C.0C04 C.0C16 0.0008 0.0028 0.0016 0.0166 0.0008 0.0008 0.0024 C.0409 0.002u C.0004 C.0004 C.0004 C.0004 C.0004 C.0004 C.0004 C.0C36 C.0C16 C.0C04 O.C053 C.0135 0.0000 0.0016 0.0024 0.0024 0.000C 0.3324 C.0020 0.0004 0.0016 0.0020 0.0097 0.0324 0.0563 C. OCC4 U. OOO8 U. OOO4 C. OO4 U. OOO4 c.103c c.ccc4 c.ccc4 3.0355 3.0012 3.0003 0.0004 0.0004 0.0006 0.0632 0.1567 0.0000 0.0034 0.0016 0.0004 0.0004 0.0004 0.0004 0.0037 0.0267 0.0012 0.0064 0.0004 0.0008 1.0136 0.0211 77 W3 M Ħ 3 Subtest Subtest Subtest

Subtest

Subtest

TABLE A-3-3 (Continued): Level 13,

0.0012 0.0012 0.0006

0.0008 0.0036 0.0045 0.0486

0.0008

Lo. K

APPENDIX IV

Frequency Distribution of Items for Each of the Eleven
Subtests with Respect to the Probability Resultant from
the Chi-Square Test of the Goodness of Fit, and the
Percentage Correct Response for Each of Those Items
Whose Resultant Probability Is 0.001 or Greater

5

TABLE A-4-1

Frequency Distribution of Items for Each of the Eleven Subtests with Respect to the Probability Resultant from the Chi-Square Test of the Goodness of Fit. The Uniform Distribution Is Assumed for the Theoretical Frequency Distribution for the Incorrect Alternatives. Number of Degrees of Freedom for Each Item Is 2, Except for the Items of Subtest L1 for Which the Number of Degrees of Freedom Is 3. (Original Data)

Level 11

	1			Pr	obabili	ty					
Subtest	.0000-	.0005-		.0055- .0105			.1005- .2005	. 2005- . 4005	.4005- .8005	.8005- 1.0000	Total
v	42					1					43
R	68	1	2			2				1	74
L.1	43										43
1.2	40										40
L3	40										40
14	32									}	32
WL	31		ı		1		1		1	1	36
W2	24	ı			1					ļ	26
W3	53						2		1	- 1	56
HL	36	1					1	1	2	1	42
H2	29										29
Total	438	3	3	0	2	3	4	1	4	3	461

Level 12

				P	robebil	167					
Subtest	.0000- .0005	.0005- .0015					.1005- .2005			.8005- 1.0000	Total
٧	42	1		1	1			1			46
R	71	2		1	1				1		76
Ll	46									1	46
1.2	42										42
L3	42										42
L4	32									j	32
W1	37	1			1			1			40
W2	24	1	2					1			28
W3	57	ı						1		- 1	59
H 1	42				1		1		1		45
H2	30			ı						1	31
Total	465		2	3	4	0	1	4	2	0	487

Level 13

				Pr	obab111	ty					
Subtest	.0000- .0005	.0005- .0015	.0015- .0055	.0055- .0105		.0505- .1005	.1005- .2005	.2005- .4005	.4005- .8005	.8005- 1.0000	Total
٧	45					1	ı			ı	48
R	71	1	1		2	1		1	ı	•	78
Li	48										48
נו	43										43
L3	43									l	43
и	32										32
٧ı	40					1					41
W2	25	ı				1				1	28
W3	57					1	1				59
MI	44	1		1			1	ı		i	48
H2	31				1					I	32
Total	479)	1	1	3	5	,	2	1	2	500

TABLE A-4-2

Percentage Correct Response for Each of the Items Whose Probability Obtained by the Chi-Square Test of the Goodness of Fit Is 0.001 or Greater. (Original Data)

Level 11

	Subtest	Item Number	Probability	Percentage Correct
,	**	61	0.061	35.9
1	V	61	0.061 0.001**	93.1
2	R	26		
2 3 4	R	42	0.081	92.9
	R	73	0.002**	64.8
5 6	R	78	0.889	79.2
6	R	88	0.004**	58.2
7 8	R	47	0.063	81.1
8	W1	13	0.005**	93.7
9	W1	20	0.036*	90.7
10	W1	21	0.824	83.9
11	W1	25	0.514	57.9
12	Wl	29	0.147	52.1
13	W2	29	0.012*	85.7
14	W2	41	0.001**	63.7
15	W3	39	0.565	63.7
16	W3	40	0.111	76.8
17	W3	53	0.142	63.3
18	M1	31	0.445	91.3
19	M1	37	0.121	53.2
20	Ml	38	0.669	80.3
21	M1	45	0.246	68.0
22	M1	63	0.001**	29.9
23	M1	68	0.815	71.7
4 1				1

TABLE A-4-2 (Continued): Level 12.

	Subtest	Item Number	Probability	Percentage Correct
1	v	49	0.009**	65.4
1 2 3	v	52	0.003***	57.0
3	v	57	0.048*	63.3
4	V	79	0.209	46.6
5	R	78	0.584	89.6
6	R	88	0.009**	73.9
7	R	105	0.001**	63.5
7 8	R	127	0.001**	53.8
9	R	131	0.018*	54.4
10	W1	28	0.001**	69.4
11	W1	29	0.289	73.8
12	W1	53	0.022*	46.0
13	W2	44	0.261	85.2
14	W2	45	0.001**	72.2
15	W2	48	0.005**	60.5
16	W2	59	0.004**	55.9
17	W3	50	0.001**	75.8
18	W3	53	0.383 İ	79.1
19	M1	68	0.016*	87.4
20	M1	74	0.171	79.3
21	Ml	95	0.660	29.4
22	M2	59	0.007**	47.2

TABLE A-4-2 (Continued): Level 13.

	Subtest	Item Number	Probability	Percentage Correct
1	V	61	0.122	57.6
2	V	79	0.063	56.6
2 3	v	93	0.851	49.9
4	R	88	0.001**	80.1
5	R	92	0.032*	86.9
6	R	103	0.002**	71.1
7	R	131	0.098	63.2
8	R	140	0.403	33.8
9	R	142	0.021*	38.0
10	R	155	0.249	35.7
11	W1	78	0.066	35.7
12	W2	44	0.915	88.1
13	W2	45	0.001**	79.7
14	W2	48	0.096	66.1
15	W 3	70	0.089	85.3
16	w3	83	0.186	50.5
17	M1	74	0.010*	87.2
18	M1	88	0.001**	73.6
19	Ml	105	0.134	71.9
20	M1	107	0.280	40.6
21	M2	75	0.050	29.4

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